Part II: Sparse Gaussian Processes

Daniel Hernández-Lobato

Computer Science Department Universidad Autónoma de Madrid

http://dhnzl.org, daniel.hernandez@uam.es

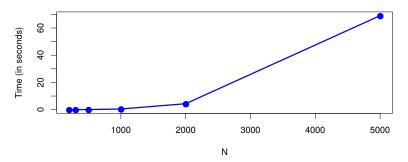
The memory cost is in $\mathcal{O}(N^2)$ since we have to compute Σ .

The memory cost is in $\mathcal{O}(N^2)$ since we have to compute Σ .

The computational cost is in $\mathcal{O}(N^3)$ since we have to invert Σ .

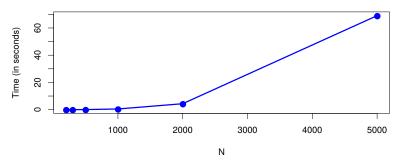
The memory cost is in $\mathcal{O}(N^2)$ since we have to compute Σ .

The computational cost is in $\mathcal{O}(N^3)$ since we have to invert Σ .



The memory cost is in $\mathcal{O}(N^2)$ since we have to compute Σ .

The computational cost is in $\mathcal{O}(N^3)$ since we have to invert Σ .



We can handle just a few thousand data instances at most!

Run the first cells of the notebook to measure GP fitting time and complete task 1!

GPs are non-parametric models whose flexibility grows with N!

GPs are non-parametric models whose flexibility grows with N!

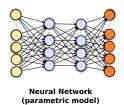
GPs are the limiting case $(H \to \infty)$ of Bayesian Neural Networks!

GPs are non-parametric models whose flexibility grows with N!

GPs are the limiting case $(H \to \infty)$ of Bayesian Neural Networks!

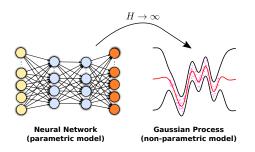
GPs are non-parametric models whose flexibility grows with N!

GPs are the limiting case $(H \to \infty)$ of Bayesian Neural Networks!



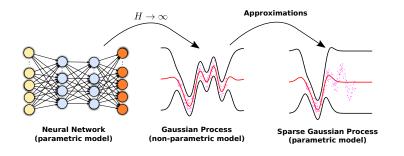
GPs are non-parametric models whose flexibility grows with N!

GPs are the limiting case $(H \to \infty)$ of Bayesian Neural Networks!



GPs are non-parametric models whose flexibility grows with N!

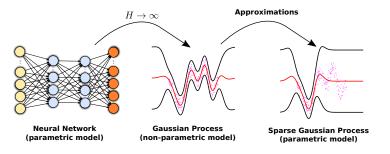
GPs are the limiting case $(H \to \infty)$ of Bayesian Neural Networks!



GPs are non-parametric models whose flexibility grows with N!

GPs are the limiting case $(H \to \infty)$ of Bayesian Neural Networks!

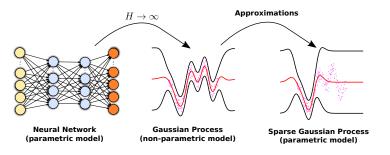
Idea: go back to the parametric model, but in such a way that we can still make inference easily!



• Nyström, Random Features and FITC: approximate GP prior!

GPs are non-parametric models whose flexibility grows with N!

GPs are the limiting case $(H \to \infty)$ of Bayesian Neural Networks!



- Nyström, Random Features and FITC: approximate GP prior!
- **VFE**: does approximate inference with a simplified distribution q.

Motivation: The posterior mean and covariances require $(I\sigma^2 + \Sigma)^{-1}$.

Motivation: The posterior mean and covariances require $(I\sigma^2 + \Sigma)^{-1}$.

Can we approximate the inverse of $I\sigma^2 + \Sigma$ with a cheaper cost?

Motivation: The posterior mean and covariances require $(I\sigma^2 + \Sigma)^{-1}$.

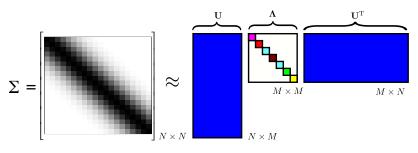
Can we approximate the inverse of $I\sigma^2 + \Sigma$ with a cheaper cost?

A low rank m approximation of Σ does the job:

Motivation: The posterior mean and covariances require $(I\sigma^2 + \Sigma)^{-1}$.

Can we approximate the inverse of $I\sigma^2 + \Sigma$ with a cheaper cost?

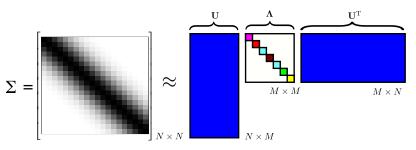
A low rank m approximation of Σ does the job:



Motivation: The posterior mean and covariances require $(I\sigma^2 + \Sigma)^{-1}$.

Can we approximate the inverse of $I\sigma^2 + \Sigma$ with a cheaper cost?

A low rank m approximation of Σ does the job:



The Woodbury formula gives $(I\sigma^2 + U\Lambda U^T)^{-1}$ with cost $\mathcal{O}(M^2N)$!

$$(A + PCQ)^{-1} = A^{-1} - A^{-1}P(C^{-1} + QA^{-1}P)^{-1}QA^{-1}$$

$$(\mathbf{A} + \mathbf{PCQ})^{-1} = \mathbf{A}^{-1} - \mathbf{A}^{-1} \mathbf{P} \left(\mathbf{C}^{-1} + \mathbf{Q} \mathbf{A}^{-1} \mathbf{P} \right)^{-1} \mathbf{Q} \mathbf{A}^{-1}$$

Let us now use $\mathbf{A} = \mathbf{I}\sigma^2$, $\mathbf{P} = \mathbf{U}$, $\mathbf{Q} = \mathbf{U}^\mathsf{T}$ and $\mathbf{C} = \mathbf{\Lambda}$.

$$(A + PCQ)^{-1} = A^{-1} - A^{-1}P(C^{-1} + QA^{-1}P)^{-1}QA^{-1}$$

Let us now use $\mathbf{A} = \mathbf{I}\sigma^2$, $\mathbf{P} = \mathbf{U}$, $\mathbf{Q} = \mathbf{U}^\mathsf{T}$ and $\mathbf{C} = \mathbf{\Lambda}$.

Note that A and C are diagonal with sizes $N \times N$ and $M \times M!$

$$(A + PCQ)^{-1} = A^{-1} - A^{-1}P(C^{-1} + QA^{-1}P)^{-1}QA^{-1}$$

Let us now use $\mathbf{A} = \mathbf{I}\sigma^2$, $\mathbf{P} = \mathbf{U}$, $\mathbf{Q} = \mathbf{U}^\mathsf{T}$ and $\mathbf{C} = \mathbf{\Lambda}$.

Note that A and C are diagonal with sizes $N \times N$ and $M \times M!$

$$\mathbf{C}^{-1} + \mathbf{Q}\mathbf{A}^{-1}\mathbf{P} = \mathbf{\Lambda}^{-1} + \mathbf{U}^{\mathsf{T}}\mathbf{U}\sigma^{-2}$$
 has size $M \times M!$

$$(A + PCQ)^{-1} = A^{-1} - A^{-1}P(C^{-1} + QA^{-1}P)^{-1}QA^{-1}$$

Let us now use $\mathbf{A} = \mathbf{I}\sigma^2$, $\mathbf{P} = \mathbf{U}$, $\mathbf{Q} = \mathbf{U}^\mathsf{T}$ and $\mathbf{C} = \mathbf{\Lambda}$.

Note that A and C are diagonal with sizes $N \times N$ and $M \times M!$

$$\mathbf{C}^{-1} + \mathbf{Q}\mathbf{A}^{-1}\mathbf{P} = \mathbf{\Lambda}^{-1} + \mathbf{U}^{\mathsf{T}}\mathbf{U}\sigma^{-2}$$
 has size $M \times M!$

$$(\mathbf{I}\sigma^2 + \mathbf{U}\Lambda\mathbf{U}^\mathsf{T})^{-1} = \mathbf{I}\sigma^{-2} - \sigma^{-2}\mathbf{U}\left(\Lambda^{-1} + \mathbf{U}^\mathsf{T}\sigma^{-2}\mathbf{U}\right)^{-1}\mathbf{U}^\mathsf{T}\sigma^{-2}$$

$$(A + PCQ)^{-1} = A^{-1} - A^{-1}P(C^{-1} + QA^{-1}P)^{-1}QA^{-1}$$

Let us now use $\mathbf{A} = \mathbf{I}\sigma^2$, $\mathbf{P} = \mathbf{U}$, $\mathbf{Q} = \mathbf{U}^\mathsf{T}$ and $\mathbf{C} = \mathbf{\Lambda}$.

Note that A and C are diagonal with sizes $N \times N$ and $M \times M!$

$$\mathbf{C}^{-1} + \mathbf{Q}\mathbf{A}^{-1}\mathbf{P} = \mathbf{\Lambda}^{-1} + \mathbf{U}^{\mathsf{T}}\mathbf{U}\sigma^{-2}$$
 has size $M \times M!$

$$(\mathbf{I}\sigma^2 + \mathbf{U}\Lambda\mathbf{U}^\mathsf{T})^{-1} = \mathbf{I}\sigma^{-2} - \sigma^{-2}\mathbf{U}\left(\Lambda^{-1} + \mathbf{U}^\mathsf{T}\sigma^{-2}\mathbf{U}\right)^{-1}\mathbf{U}^\mathsf{T}\sigma^{-2}$$

Computing the whole expression has cost $\mathcal{O}(NM^2)$!

Eigenfunction Analysis of Covariance Functions

GPs are equivalent to a Bayesian linear model on an extended input space given by the eigenfunctions of the covariance function.

Eigenfunction Analysis of Covariance Functions

GPs are equivalent to a Bayesian linear model on an extended input space given by the eigenfunctions of the covariance function.

Extended input space: A function $\phi(\cdot)$ that obeys

$$\int C(\mathbf{x}, \mathbf{x}') \phi(\mathbf{x}) p(\mathbf{x}) d\mathbf{x} = \lambda \phi(\mathbf{x}'),$$

is an eigenfunction of $C(\cdot, \cdot)$ with eigenvalue λ , w.r.t., $p(\mathbf{x})$.

Eigenfunction Analysis of Covariance Functions

GPs are equivalent to a Bayesian linear model on an extended input space given by the eigenfunctions of the covariance function.

Extended input space: A function $\phi(\cdot)$ that obeys

$$\int C(\mathbf{x}, \mathbf{x}') \phi(\mathbf{x}) p(\mathbf{x}) d\mathbf{x} = \lambda \phi(\mathbf{x}'),$$

is an eigenfunction of $C(\cdot,\cdot)$ with eigenvalue λ , w.r.t., $p(\mathbf{x})$.

Mercer's theorem:

$$C(\mathbf{x}, \mathbf{x}') = \sum_{i=1}^{\infty} \lambda_i \phi_i(\mathbf{x}) \phi_i(\mathbf{x}').$$

An Analytic Example

Consider:

$$p(x) = \mathcal{N}(x|0,\sigma^2), \qquad C(x,x') = \exp\left\{-\frac{1}{2\ell^2}(x-x')^2\right\}.$$

An Analytic Example

Consider:

$$p(x) = \mathcal{N}(x|0,\sigma^2), \qquad C(x,x') = \exp\left\{-\frac{1}{2\ell^2}(x-x')^2\right\}.$$

Then,

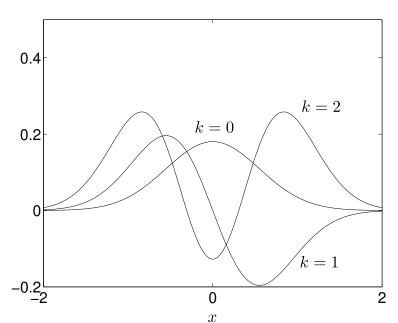
$$\lambda_k = \sqrt{\frac{2a}{A}}B^k$$
, $\phi_k(x) = \exp\left\{-(c-a)x^2\right\}H_k(\sqrt{2c}x)$,

for k = 0, 1, 2, ..., with

$$a^{-1} = 4\sigma^2$$
, $b^{-1} = 2\ell^2$, $c = \sqrt{a^2 + 2ab}$, $A = a + b + c$, $B = b/a$,

and $H_k(\cdot)$, the k-th order Hermite polynomial.

Hermite Polynomials

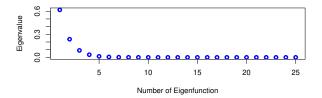


Covariance Function Approximation

Considering only the first eigenfunctions and eigenvalues is expected to give a good approximation of the covariance function!

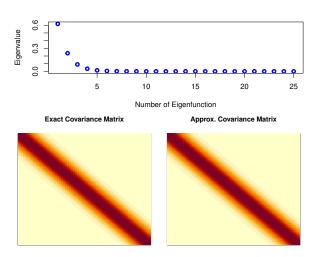
Covariance Function Approximation

Considering only the first eigenfunctions and eigenvalues is expected to give a good approximation of the covariance function!



Covariance Function Approximation

Considering only the first eigenfunctions and eigenvalues is expected to give a good approximation of the covariance function!



Nyström Approximation of Eigenfunctions

Let p(x) be the distribution of the observed data.

Nyström Approximation of Eigenfunctions

Let p(x) be the distribution of the observed data.

Consider the Monte Carlo estimator:

$$\lambda_i \phi_i(\mathbf{x}') = \int C(\mathbf{x}, \mathbf{x}') \phi_i(\mathbf{x}) p(\mathbf{x}) d\mathbf{x} \approx \frac{1}{N} \sum_{n=1}^N C(\mathbf{x}_n, \mathbf{x}') \phi_i(\mathbf{x}_n).$$

Nyström Approximation of Eigenfunctions

Let p(x) be the distribution of the observed data.

Consider the Monte Carlo estimator:

$$\lambda_i \phi_i(\mathbf{x}') = \int C(\mathbf{x}, \mathbf{x}') \phi_i(\mathbf{x}) p(\mathbf{x}) d\mathbf{x} \approx \frac{1}{N} \sum_{n=1}^N C(\mathbf{x}_n, \mathbf{x}') \phi_i(\mathbf{x}_n).$$

For \mathbf{x}' in the training set, this motivates the following eigenvalue problem:

$$\lambda_i^{\mathsf{mat}} \mathbf{u}_i = \mathbf{\Sigma} \mathbf{u}_i$$

with $\Sigma_{i,j} = C(\mathbf{x}_i, \mathbf{x}_j)$. Then, we approximate $\phi_i(\mathbf{x}_j) \approx \sqrt{N}(\mathbf{u}_i)_j = \tilde{\phi}_i(\mathbf{x}_j)$, and $\lambda_i \approx \lambda_i^{\text{mat}}/N = \tilde{\lambda}_i$, which guarantees that $\Sigma = \tilde{\Phi} \tilde{\Lambda} \tilde{\Phi}^{\mathsf{T}}$, with $j = 1, \dots, N$.

Nyström Approximation of Eigenfunctions

Let p(x) be the distribution of the observed data.

Consider the Monte Carlo estimator:

$$\lambda_i \phi_i(\mathbf{x}') = \int C(\mathbf{x}, \mathbf{x}') \phi_i(\mathbf{x}) p(\mathbf{x}) d\mathbf{x} \approx \frac{1}{N} \sum_{n=1}^N C(\mathbf{x}_n, \mathbf{x}') \phi_i(\mathbf{x}_n).$$

For \mathbf{x}' in the training set, this motivates the following eigenvalue problem:

$$\lambda_i^{\mathsf{mat}} \mathsf{u}_i = \mathbf{\Sigma} \mathsf{u}_i \,,$$

with $\Sigma_{i,j} = C(\mathbf{x}_i, \mathbf{x}_j)$. Then, we approximate $\phi_i(\mathbf{x}_j) \approx \sqrt{N}(\mathbf{u}_i)_j = \tilde{\phi}_i(\mathbf{x}_j)$, and $\lambda_i \approx \lambda_i^{\mathrm{mat}}/N = \tilde{\lambda}_i$, which guarantees that $\Sigma = \tilde{\Phi} \tilde{\Lambda} \tilde{\Phi}^{\mathsf{T}}$, with $j = 1, \ldots, N$.

For an arbitrary \mathbf{x}' not in the training set, then:

$$\tilde{\phi}_i(\mathbf{x}') = \frac{1}{N\lambda_i} \sum_{n=1}^N C(\mathbf{x}', \mathbf{x}_n) \phi_i(\mathbf{x}_n) \approx \frac{\sqrt{N}}{\lambda_i^{\mathsf{mat}}} \sum_{n=1}^N C(\mathbf{x}', \mathbf{x}_n) (\mathbf{u}_i)_n = \frac{\sqrt{N}}{\lambda_i^{\mathsf{mat}}} \mathbf{\Sigma}(\mathbf{x}')^{\mathsf{T}} \mathbf{u}_i.$$

We choose a random subset of size M < N of the training data, to approximate the eigenfunctions and eigenvalues!

We choose a random subset of size M < N of the training data, to approximate the eigenfunctions and eigenvalues!

Using Mercer's theorem and the previous approximation, we approximate the covariance function as:

$$C(\mathbf{x},\mathbf{x}') = \sum_{i=1}^{\infty} \lambda_i \phi_i(\mathbf{x}) \phi_i(\mathbf{x}') \approx \sum_{i=1}^{M} \tilde{\lambda}_i \tilde{\phi}_i(\mathbf{x}) \tilde{\phi}_i(\mathbf{x}').$$

We choose a random subset of size M < N of the training data, to approximate the eigenfunctions and eigenvalues!

Using Mercer's theorem and the previous approximation, we approximate the covariance function as:

$$C(\mathbf{x},\mathbf{x}') = \sum_{i=1}^{\infty} \lambda_i \phi_i(\mathbf{x}) \phi_i(\mathbf{x}') \approx \sum_{i=1}^{M} \tilde{\lambda}_i \tilde{\phi}_i(\mathbf{x}) \tilde{\phi}_i(\mathbf{x}').$$

which results in a rank M approximation of the covariance matrix Σ :

$$oldsymbol{\Sigma} pprox ilde{oldsymbol{\Sigma}} = ilde{oldsymbol{\Phi}} ilde{oldsymbol{\Phi}}^{\mathsf{T}} = oldsymbol{\Sigma}_{\mathcal{N},\mathcal{M}} oldsymbol{\Sigma}_{\mathcal{M},\mathcal{M}}^{-1} oldsymbol{\Sigma}_{\mathcal{M},\mathcal{N}} \,.$$

We choose a random subset of size M < N of the training data, to approximate the eigenfunctions and eigenvalues!

Using Mercer's theorem and the previous approximation, we approximate the covariance function as:

$$C(\mathbf{x},\mathbf{x}') = \sum_{i=1}^{\infty} \lambda_i \phi_i(\mathbf{x}) \phi_i(\mathbf{x}') \approx \sum_{i=1}^{M} \tilde{\lambda}_i \tilde{\phi}_i(\mathbf{x}) \tilde{\phi}_i(\mathbf{x}').$$

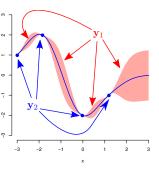
which results in a rank M approximation of the covariance matrix Σ :

$$oldsymbol{\Sigma} pprox ilde{oldsymbol{\Sigma}} = ilde{oldsymbol{\Phi}} ilde{oldsymbol{\Phi}}^{\mathsf{T}} = oldsymbol{\Sigma}_{\mathcal{N},\mathcal{M}} oldsymbol{\Sigma}_{\mathcal{M},\mathcal{M}}^{-1} oldsymbol{\Sigma}_{\mathcal{M},\mathcal{N}} \,.$$

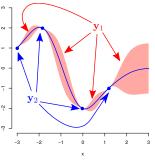
The inverse of $I\sigma^2 + \tilde{\Sigma}$ can be efficiently computed using the Woodbury formula with cost $\mathcal{O}(NM^2)!$

$$\rho(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c}\mathbf{f}\\\mathbf{f}^{\star}\end{array}\right] \middle| \left[\begin{array}{c}\mathbf{0}\\\mathbf{0}\end{array}\right], \left[\begin{array}{cc}\mathbf{\tilde{\Sigma}} & \mathbf{\Sigma}_{\mathbf{f}^{\star}}\\\mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}} & \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}}\end{array}\right]\right)$$

$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \left[\begin{array}{cc} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \mathbf{\tilde{\Sigma}} & \mathbf{\Sigma}_{\mathbf{f}^{\star}} \\ \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}} & \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array}\right]\right)$$

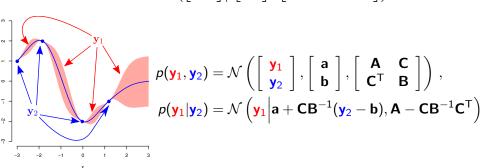


$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{cc} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \mathbf{\tilde{\Sigma}} & \mathbf{\Sigma}_{\mathbf{f}\mathbf{f}^{\star}} \\ \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}} & \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array}\right]\right)$$



$$\rho(\mathbf{y_1}, \mathbf{y_2}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{y_1} \\ \mathbf{y_2} \end{array}\right], \left[\begin{array}{c} \mathbf{a} \\ \mathbf{b} \end{array}\right], \left[\begin{array}{c} \mathbf{A} & \mathbf{C} \\ \mathbf{C}^\mathsf{T} & \mathbf{B} \end{array}\right]\right),$$

$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{cc} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \mathbf{\tilde{\Sigma}} & \mathbf{\Sigma}_{\mathbf{f}\mathbf{f}^{\star}} \\ \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}} & \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array}\right]\right)$$



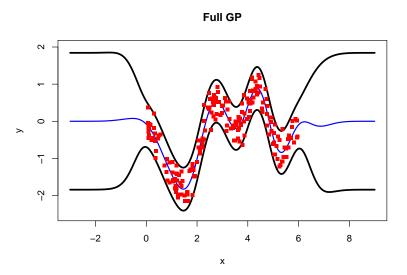
$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\begin{bmatrix} \mathbf{f} \\ \mathbf{f}^{\star} \end{bmatrix} \middle| \begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \boldsymbol{\Sigma} & \boldsymbol{\Sigma}_{\mathbf{f}^{\star}} \\ \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}} & \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{bmatrix}\right)$$

$$p(\mathbf{y}_{1}, \mathbf{y}_{2}) = \mathcal{N}\left(\begin{bmatrix} \mathbf{y}_{1} \\ \mathbf{y}_{2} \end{bmatrix}, \begin{bmatrix} \mathbf{a} \\ \mathbf{b} \end{bmatrix}, \begin{bmatrix} \mathbf{A} & \mathbf{C} \\ \mathbf{C}^{\mathsf{T}} & \mathbf{B} \end{bmatrix}\right),$$

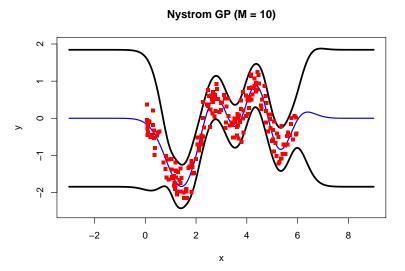
$$p(\mathbf{y}_{1}|\mathbf{y}_{2}) = \mathcal{N}\left(\mathbf{y}_{1} \middle| \mathbf{a} + \mathbf{C}\mathbf{B}^{-1}(\mathbf{y}_{2} - \mathbf{b}), \mathbf{A} - \mathbf{C}\mathbf{B}^{-1}\mathbf{C}^{\mathsf{T}}\right)$$

$$p(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}| \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}} \boldsymbol{\tilde{\Sigma}}_{\mathbf{f}^{\dagger}}^{-1} \mathbf{f}, \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}} \boldsymbol{\tilde{\Sigma}}_{\mathbf{f}^{\dagger}}^{-1} \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}}^{\mathsf{T}}\right)$$

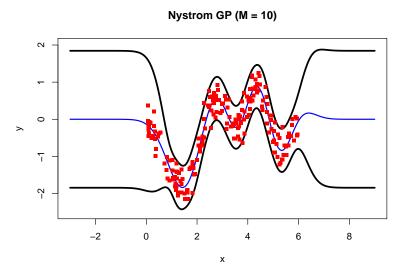
Nyström Approximation: Illustrative Example



Nyström Approximation: Illustrative Example



Nyström Approximation: Illustrative Example



The approximation is similar to the full GP in some regions!

• Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- If M = N the method is exact since $\tilde{\Sigma} = \Sigma$.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- If M=N the method is exact since $\tilde{\Sigma}=\Sigma$.
- For small *M* it can give bad results according to empirical evidence.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- If M=N the method is exact since $\tilde{\Sigma}=\Sigma$.
- For small M it can give bad results according to empirical evidence.
- ullet It can perform well if Σ is dominated by a few eigenvalues.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- If M=N the method is exact since $\tilde{\Sigma}=\Sigma$.
- For small *M* it can give bad results according to empirical evidence.
- It can perform well if Σ is dominated by a few eigenvalues.
- As the *M* points are chosen at random it may give different results.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- If M=N the method is exact since $\tilde{\Sigma}=\Sigma$.
- For small *M* it can give bad results according to empirical evidence.
- It can perform well if Σ is dominated by a few eigenvalues.
- As the *M* points are chosen at random it may give different results.
- Since the approximation is done only over the covariance matrix of the training data, negative predictive variances are possible, but rare.

Run the cells of the notebook to fit a sparse GP using the Nyström approximation and complete task 2!

Run the cells of the notebook compare the exact and approximate predictive distribution and complete task 3!

They can be used to approximate any stationary covariance function (only depends on the distance between points).

They can be used to approximate any stationary covariance function (only depends on the distance between points).

Bochner's theorem:

A covariance function $C(\mathbf{x}, \mathbf{x}') = C(\mathbf{x} - \mathbf{x}')$ on \mathbb{R}^D is positive definite if and only if $C(\mathbf{x} - \mathbf{x}')$ is the Fourier transform of a distribution $s(\mathbf{w})$.

They can be used to approximate any stationary covariance function (only depends on the distance between points).

Bochner's theorem:

A covariance function $C(\mathbf{x}, \mathbf{x}') = C(\mathbf{x} - \mathbf{x}')$ on \mathbb{R}^D is positive definite if and only if $C(\mathbf{x} - \mathbf{x}')$ is the Fourier transform of a distribution $s(\mathbf{w})$.

$$C(\mathbf{x}, \mathbf{x}') = \int \exp\{-i\mathbf{w}^{\mathsf{T}}(\mathbf{x} - \mathbf{x}')\}s(\mathbf{w})d\mathbf{w},$$

$$s(\mathbf{w}) = \frac{1}{(2\pi)^{D}}\int \exp\{i\mathbf{w}^{\mathsf{T}}\boldsymbol{\tau}\}C(\boldsymbol{\tau}, \mathbf{0})d\boldsymbol{\tau}.$$

They can be used to approximate any stationary covariance function (only depends on the distance between points).

Bochner's theorem:

A covariance function $C(\mathbf{x}, \mathbf{x}') = C(\mathbf{x} - \mathbf{x}')$ on \mathbb{R}^D is positive definite if and only if $C(\mathbf{x} - \mathbf{x}')$ is the Fourier transform of a distribution $s(\mathbf{w})$.

$$egin{aligned} \mathcal{C}(\mathbf{x},\mathbf{x}') &= \int \exp\{-i\mathbf{w}^\mathsf{T}(\mathbf{x}-\mathbf{x}')\}s(\mathbf{w})d\mathbf{w}\,, \ s(\mathbf{w}) &= rac{1}{(2\pi)^D}\int \exp\{i\mathbf{w}^\mathsf{T}oldsymbol{ au}\}\mathcal{C}(oldsymbol{ au},oldsymbol{0})doldsymbol{ au}\,. \end{aligned}$$

 $s(\mathbf{w})$ is called the spectral density of the covariance function.

Due to Bochner's theorem, the covariance can be written as:

$$\begin{split} C(\mathbf{x}, \mathbf{x}') &= \mathbb{E}_{s(\mathbf{w})} \left[\exp\{-i\mathbf{w}^\mathsf{T} (\mathbf{x} - \mathbf{x}')\} \right] \\ &= 2 \mathbb{E}_{s(\mathbf{w}), b \sim U[0, 2\pi]} \left[\cos\left(\mathbf{w}^\mathsf{T} \mathbf{x} + b\right) \cos\left(\mathbf{w}^\mathsf{T} \mathbf{x}' + b\right) \right] \,. \end{split}$$

Due to Bochner's theorem, the covariance can be written as:

$$C(\mathbf{x}, \mathbf{x}') = \mathbb{E}_{s(\mathbf{w})} \left[\exp\{-i\mathbf{w}^{\mathsf{T}}(\mathbf{x} - \mathbf{x}')\} \right]$$
$$= 2\mathbb{E}_{s(\mathbf{w}), b \sim U[0, 2\pi]} \left[\cos\left(\mathbf{w}^{\mathsf{T}}\mathbf{x} + b\right) \cos\left(\mathbf{w}^{\mathsf{T}}\mathbf{x}' + b\right) \right].$$

The expectation can be approximated by a Monte Carlo average!

Due to Bochner's theorem, the covariance can be written as:

$$C(\mathbf{x}, \mathbf{x}') = \mathbb{E}_{s(\mathbf{w})} \left[\exp\{-i\mathbf{w}^{\mathsf{T}}(\mathbf{x} - \mathbf{x}')\} \right]$$
$$= 2\mathbb{E}_{s(\mathbf{w}), b \sim U[0, 2\pi]} \left[\cos\left(\mathbf{w}^{\mathsf{T}}\mathbf{x} + b\right) \cos\left(\mathbf{w}^{\mathsf{T}}\mathbf{x}' + b\right) \right].$$

The expectation can be approximated by a Monte Carlo average!

We can reduce the variance of the estimator by generating M samples:

$$C(\mathbf{x}, \mathbf{x}') \approx \frac{2}{M} \sum_{m=1}^{M} \cos\left(\mathbf{w}_{m}^{\mathsf{T}} \mathbf{x} + b_{m}\right) \cos\left(\mathbf{w}_{m}^{\mathsf{T}} \mathbf{x}' + b_{m}\right) = \phi(\mathbf{x})^{\mathsf{T}} \phi(\mathbf{x}')$$

with $\phi(\mathbf{x}) = \sqrt{\frac{2}{M}} \cos \left(\mathbf{W}^\mathsf{T} \mathbf{x} + \mathbf{b} \right)$ a random M feature expansion.

Due to Bochner's theorem, the covariance can be written as:

$$C(\mathbf{x}, \mathbf{x}') = \mathbb{E}_{s(\mathbf{w})} \left[\exp\{-i\mathbf{w}^{\mathsf{T}}(\mathbf{x} - \mathbf{x}')\} \right]$$
$$= 2\mathbb{E}_{s(\mathbf{w}), b \sim U[0, 2\pi]} \left[\cos\left(\mathbf{w}^{\mathsf{T}}\mathbf{x} + b\right) \cos\left(\mathbf{w}^{\mathsf{T}}\mathbf{x}' + b\right) \right].$$

The expectation can be approximated by a Monte Carlo average!

We can reduce the variance of the estimator by generating M samples:

$$C(\mathbf{x}, \mathbf{x}') \approx \frac{2}{M} \sum_{m=1}^{M} \cos\left(\mathbf{w}_{m}^{\mathsf{T}} \mathbf{x} + b_{m}\right) \cos\left(\mathbf{w}_{m}^{\mathsf{T}} \mathbf{x}' + b_{m}\right) = \phi(\mathbf{x})^{\mathsf{T}} \phi(\mathbf{x}')$$

with $\phi(\mathbf{x}) = \sqrt{\frac{2}{M}} \cos \left(\mathbf{W}^\mathsf{T} \mathbf{x} + \mathbf{b} \right)$ a random M feature expansion.

For the squared exponential covariance function $s(\mathbf{w})$ is Gaussian!

Approximate Covariance Function

The covariance matrix can be simply approximated as:

$$oldsymbol{\Sigma}pprox ilde{oldsymbol{\Sigma}}=oldsymbol{\Phi}oldsymbol{\Phi}^\mathsf{T}$$

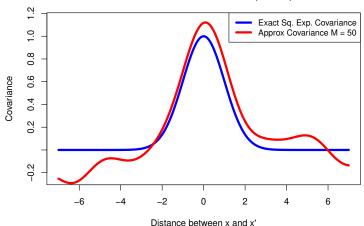
and hence $\mathbf{I}\sigma^2 + \tilde{\Sigma}$ can be inverted with cost $\mathcal{O}(NM^2)$.

Approximate Covariance Function

The covariance matrix can be simply approximated as:

$$\mathbf{\Sigma} pprox \mathbf{ ilde{\Sigma}} = \mathbf{\Phi} \mathbf{\Phi}^\mathsf{T}$$

and hence $\mathbf{I}\sigma^2 + \tilde{\Sigma}$ can be inverted with cost $\mathcal{O}(NM^2)$.



$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{cc} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \tilde{\boldsymbol{\Sigma}} & \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}} \\ \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}} & \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array}\right]\right)$$

We want to compute the value of f^* at a new \mathbf{x}^* :

$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{cc} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \tilde{\boldsymbol{\Sigma}} & \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}} \\ \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}} & \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array}\right]\right)$$

All prior covariances are now approximated using dot products with the random features computed before!

We want to compute the value of f^* at a new \mathbf{x}^* :

$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{cc} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \tilde{\boldsymbol{\Sigma}} & \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}} \\ \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}} & \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array}\right]\right)$$

All prior covariances are now approximated using dot products with the random features computed before!

$$\rho(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}|\ \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{f}, \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}}^{\mathsf{T}}\right)$$

Predictive Distribution

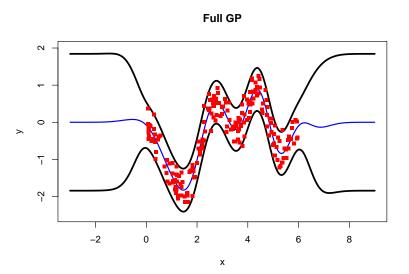
We want to compute the value of f^* at a new \mathbf{x}^* :

$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{cc} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \tilde{\boldsymbol{\Sigma}} & \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}} \\ \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}} & \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array}\right]\right)$$

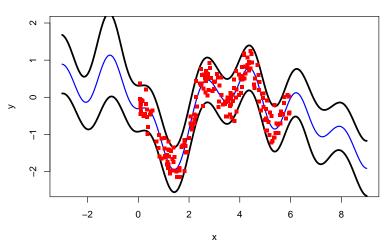
All prior covariances are now approximated using dot products with the random features computed before!

$$p(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}|\ \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{f}, \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\star}\mathbf{f}}^{\mathsf{T}}\right)$$

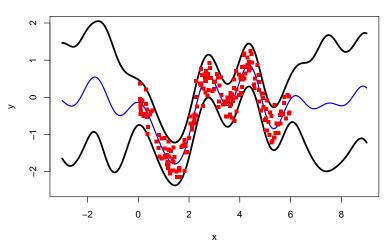
The computational cost is $\mathcal{O}(NM^2)$!



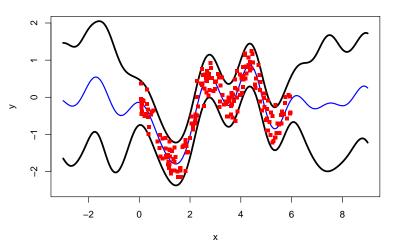












In regions with no data the approximation may be wiggling a lot!

• Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- For small *M* it can give bad results due to the wiggling effect of cosine features.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- For small M it can give bad results due to the wiggling effect of cosine features.
- Guaranteed to be exact only for $M \to \infty$.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- For small M it can give bad results due to the wiggling effect of cosine features.
- Guaranteed to be exact only for $M \to \infty$.
- It is restricted to stationary covariance functions.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- For small M it can give bad results due to the wiggling effect of cosine features.
- Guaranteed to be exact only for $M \to \infty$.
- It is restricted to stationary covariance functions.
- Very simple to implement!

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- For small M it can give bad results due to the wiggling effect of cosine features.
- Guaranteed to be exact only for $M \to \infty$.
- It is restricted to stationary covariance functions.
- Very simple to implement!
- Equivalent to a neural network with a hidden layer with M units and cosine activations, and a Bayesian linear model in the last layer!

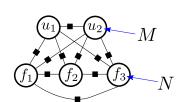
Run the cells of the notebook to fit a sparse GP using the Random Features approximation and complete task 4!

1. Extend model with $M \ll N$ inducing points and outputs at $\overline{\mathbf{X}}$.

$$\textit{p}(\mathbf{f},\mathbf{u}) = \mathcal{N}\left(\left[\begin{array}{c}\mathbf{f}\\\mathbf{u}\end{array}\right] \middle| \left[\begin{array}{c}\mathbf{0}\\\mathbf{0}\end{array}\right], \left[\begin{array}{cc}\boldsymbol{\Sigma}_{\mathbf{ff}} & \boldsymbol{\Sigma}_{\mathbf{fu}}\\\boldsymbol{\Sigma}_{\mathbf{uf}} & \boldsymbol{\Sigma}_{\mathbf{uu}}\end{array}\right]\right)$$

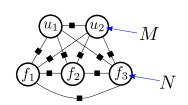
1. Extend model with $M \ll N$ inducing points and outputs at $\overline{\mathbf{X}}$.

$$\label{eq:posterior} \textit{p}(f,u) = \mathcal{N}\left(\left[\begin{array}{c} f \\ u \end{array}\right] \middle| \left[\begin{array}{c} 0 \\ 0 \end{array}\right], \left[\begin{array}{cc} \Sigma_{ff} & \Sigma_{fu} \\ \Sigma_{uf} & \Sigma_{uu} \end{array}\right]\right)$$



1. Extend model with $M \ll N$ inducing points and outputs at $\overline{\mathbf{X}}$.

$$\label{eq:posterior} \rho(f,u) = \mathcal{N}\left(\left[\begin{array}{c} f \\ u \end{array}\right] \middle| \left[\begin{array}{c} 0 \\ 0 \end{array}\right], \left[\begin{array}{cc} \Sigma_{ff} & \Sigma_{fu} \\ \Sigma_{uf} & \Sigma_{uu} \end{array}\right]\right)$$

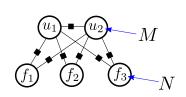


2. Introduce conditional independences:

$$p(\mathbf{f}|\mathbf{u}) = \prod_{i=1}^{N} p(f_i|\mathbf{u})$$

1. Extend model with $M \ll N$ inducing points and outputs at $\overline{\mathbf{X}}$.

$$\textit{p}(f,u) = \mathcal{N}\left(\left[\begin{array}{c} f \\ u \end{array}\right] \middle| \left[\begin{array}{c} 0 \\ 0 \end{array}\right], \left[\begin{array}{cc} \Sigma_{ff} & \Sigma_{fu} \\ \Sigma_{uf} & \Sigma_{uu} \end{array}\right]\right)$$

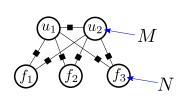


2. Introduce conditional independences:

$$p(\mathbf{f}|\mathbf{u}) = \prod_{i=1}^{N} p(f_i|\mathbf{u})$$

1. Extend model with $M \ll N$ inducing points and outputs at $\overline{\mathbf{X}}$.

$$\textit{p}(f,u) = \mathcal{N}\left(\left[\begin{array}{c} f \\ u \end{array}\right] \middle| \left[\begin{array}{c} 0 \\ 0 \end{array}\right], \left[\begin{array}{cc} \Sigma_{ff} & \Sigma_{fu} \\ \Sigma_{uf} & \Sigma_{uu} \end{array}\right]\right)$$



2. Introduce conditional independences:

$$p(\mathbf{f}|\mathbf{u}) = \prod_{i=1}^{N} p(f_i|\mathbf{u})$$

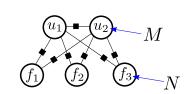
3. Marginalize \mathbf{u} to obtain an approximate GP prior for \mathbf{f} .

$$p(\mathbf{f}) = \int p(\mathbf{f}|\mathbf{u})p(\mathbf{u})d\mathbf{u} = \prod_{i=1}^{N} p(f_i|\mathbf{u})p(\mathbf{u})d\mathbf{u} = \mathcal{N}(\mathbf{f}|0, \tilde{\Sigma}_{\mathbf{f}\mathbf{f}})$$

where $\tilde{\Sigma}_{\rm ff} = D + Q_{\rm ff}$ with D diagonal and $Q_{\rm ff} = \Sigma_{\rm fu} \Sigma_{\rm uu}^{-1} \Sigma_{\rm uf}$ of rank M.

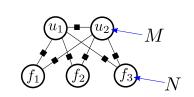
5. We make the prediction of f^* at \mathbf{x}^* by considering the approximate GP prior:

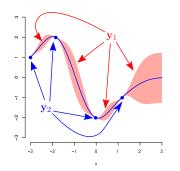
$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}\mathbf{f}} & \mathbf{Q}_{\mathbf{f}\mathbf{f}^{\star}} \\ \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}} & \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array}\right]\right)$$



5. We make the prediction of f^* at \mathbf{x}^* by considering the approximate GP prior:

$$\rho(\mathbf{f},\mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c}\mathbf{f}\\\mathbf{f}^{\star}\end{array}\right]\middle|\left[\begin{array}{c}\mathbf{0}\\\mathbf{0}\end{array}\right],\left[\begin{array}{cc}\mathbf{\tilde{\Sigma}_{ff}} & \mathbf{Q}_{ff^{\star}}\\\mathbf{Q}_{f^{\star}f} & \mathbf{\Sigma}_{f^{\star}f^{\star}}\end{array}\right]\right)$$

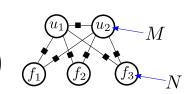




$$\begin{split} & p(\mathbf{y_1}, \mathbf{y_2}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{y_1} \\ \mathbf{y_2} \end{array} \right], \left[\begin{array}{c} \mathbf{a} \\ \mathbf{b} \end{array} \right], \left[\begin{array}{c} \mathbf{A} & \mathbf{C} \\ \mathbf{C}^\mathsf{T} & \mathbf{B} \end{array} \right] \right) \,, \\ & p(\mathbf{y_1}|\mathbf{y_2}) = \frac{p(\mathbf{y_1}, \mathbf{y_2})}{p(\mathbf{y_2})} \,, \end{split}$$

5. We make the prediction of f^* at \mathbf{x}^* by considering the approximate GP prior:

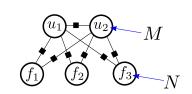
$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}} & \mathbf{Q}_{\mathbf{f}^{\star}} \\ \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}} & \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array}\right]\right)$$



$$p(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}|\ \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{f}, \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}^{\mathsf{T}}\right)$$

5. We make the prediction of f^* at \mathbf{x}^* by considering the approximate GP prior:

$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \mathbf{\tilde{\Sigma}_{ff}} & \mathbf{Q_{ff^{\star}}} \\ \mathbf{Q_{f^{\star}f}} & \mathbf{\Sigma_{f^{\star}f^{\star}}} \end{array}\right]\right)$$

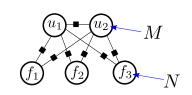


$$p(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}|\ \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{f}, \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}^{\mathsf{T}}\right)$$

Due to the structure in $\tilde{\Sigma}_{ff}$ all computations have cost in $\mathcal{O}(NM^2)$.

5. We make the prediction of f^* at \mathbf{x}^* by considering the approximate GP prior:

$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \mathbf{\tilde{\Sigma}_{ff}} & \mathbf{Q_{ff^{\star}}} \\ \mathbf{Q_{f^{\star}f}} & \mathbf{\Sigma_{f^{\star}f^{\star}}} \end{array}\right]\right)$$



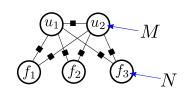
$$p(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}|\ \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{f}, \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}^{\mathsf{T}}\right)$$

Due to the structure in $\tilde{\Sigma}_{ff}$ all computations have cost in $\mathcal{O}(NM^2)$.

6. How do we find the location of the inducing points $\overline{\mathbf{X}}$?

5. We make the prediction of f^* at \mathbf{x}^* by considering the approximate GP prior:

$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array}\right] \middle| \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{c} \tilde{\mathbf{\Sigma}}_{\mathbf{f}} & \mathbf{Q}_{\mathbf{f}^{\star}} \\ \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}} & \Sigma_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array}\right]\right)$$



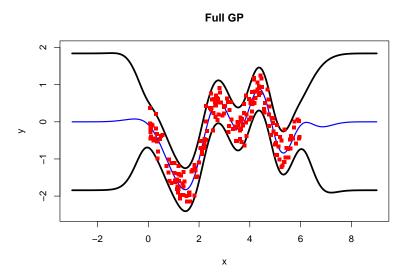
$$p(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}|\ \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{f}, \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}^{\mathsf{T}}\right)$$

Due to the structure in $\tilde{\Sigma}_{ff}$ all computations have cost in $\mathcal{O}(NM^2)$.

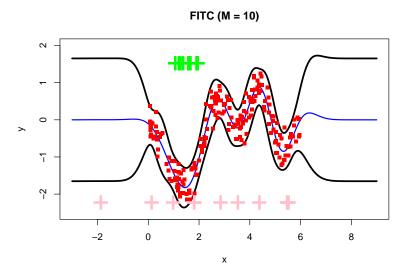
6. How do we find the location of the inducing points $\overline{\mathbf{X}}$?

Simply treat them as prior parameters and maximize the approximate marginal likelihood $p(\mathbf{f}|\mathbf{0}, \tilde{\Sigma}_{\mathbf{f}})!$

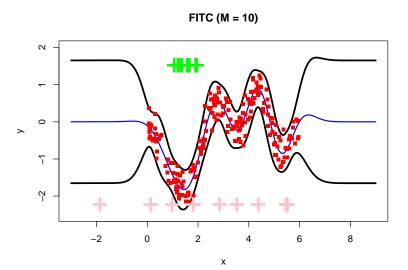
FITC: Illustrative Example



FITC: Illustrative Example



FITC: Illustrative Example



The inducing points cover the regions where the function changes!

• Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- The optimized inducing points spread over the input space where the latent function changes.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- The optimized inducing points spread over the input space where the latent function changes.
- Guaranteed to be exact if M = N and the inducing points are not optimized and located at the training points.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- The optimized inducing points spread over the input space where the latent function changes.
- Guaranteed to be exact if M = N and the inducing points are not optimized and located at the training points.
- It can be understood as considering heteroscedastic (input dependent) noise!

Run the cells of the notebook to fit a sparse GP using the FITC approximation and complete task 5!

Combines FITC with the use of Expectation Propagation to address binary classification problems!

Combines FITC with the use of Expectation Propagation to address binary classification problems!



Combines FITC with the use of Expectation Propagation to address binary classification problems!



Assumes $y_i \in \{-1, 1\}$ and a probit likelihood:

$$p(y_i|f(\mathbf{x}_i)) = \phi(y_if(\mathbf{x}_i)), \quad \phi(\cdot) \equiv \text{The c.d.f. of a standard Gaussian.}$$

Combines FITC with the use of Expectation Propagation to address binary classification problems!



Assumes $y_i \in \{-1, 1\}$ and a probit likelihood:

$$p(y_i|f(\mathbf{x}_i)) = \phi(y_if(\mathbf{x}_i)), \quad \phi(\cdot) \equiv \text{The c.d.f. of a standard Gaussian.}$$

Approximates with a Gaussian distribution the intractable posterior:

$$p(\mathbf{f}|\mathbf{y}) = \frac{\prod_{i=1}^{N} \phi(y_i f(\mathbf{x}_i)) \mathcal{N}(\mathbf{f}|\mathbf{0}, \tilde{\boldsymbol{\Sigma}})}{p(\mathbf{y})},$$

where $\tilde{\Sigma}$ is the approximate FITC covariance matrix.

Introduction to Expectation Propagation

Approximates an intractable distribution p by a parametric distribution q.

Introduction to Expectation Propagation

Approximates an intractable distribution p by a parametric distribution q.

It is based on the minimization of the KL-divergence, KL(p||q):

$$\int p(\mathbf{x}) \log \frac{p(\mathbf{x})}{q(\mathbf{x})} d\mathbf{x} = \mathsf{KL}(q|p) \geq 0.$$

Introduction to Expectation Propagation

Approximates an intractable distribution p by a parametric distribution q.

It is based on the minimization of the KL-divergence, KL(p||q):

$$\int p(\mathbf{x}) \log \frac{p(\mathbf{x})}{q(\mathbf{x})} d\mathbf{x} = \mathsf{KL}(q|p) \geq 0.$$

q is restricted to belong to a family of distributions closed under the product and ratio operation: The exponential family.

Introduction to Expectation Propagation

Approximates an intractable distribution p by a parametric distribution q.

It is based on the minimization of the KL-divergence, KL(p||q):

$$\int p(\mathbf{x}) \log \frac{p(\mathbf{x})}{q(\mathbf{x})} d\mathbf{x} = \mathsf{KL}(q|p) \geq 0.$$

q is restricted to belong to a family of distributions closed under the product and ratio operation: The exponential family.

The exponential family:

$$q(\mathbf{x}) = \exp\left(\eta^{\mathsf{T}}\mathbf{u}(\mathbf{x}) - g(\eta)\right), \quad g(\eta) = \log\int \exp\left(\eta^{\mathsf{T}}\mathbf{u}(\mathbf{x})\right)d\mathbf{x}$$

where η is a vector of natural parameters of q, $\mathbf{u}(\mathbf{x})$ are the sufficient statistics and $g(\eta)$ is a **log partition function**.

Examples of Distributions in the Exponential Family

Gaussian:

$$\mathcal{N}(x|\mu,\sigma^2) = 1/\sqrt{2\pi\sigma^2} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$$

Examples of Distributions in the Exponential Family

Gaussian:

$$\mathcal{N}(x|\mu,\sigma^2) = 1/\sqrt{2\pi\sigma^2} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$$

Exponential form:

$$\mathcal{N}(x|\mu, \sigma^2) = \exp\left(\boldsymbol{\eta}^\mathsf{T} \mathbf{u}(x) - g(\boldsymbol{\eta})\right)$$

$$\eta = (\mu/\sigma^2, 1.0/\sigma^2)^{\mathsf{T}}, \quad \mathbf{u}(x) = (x, -0.5x^2)^{\mathsf{T}}, \quad g(\eta) = -\frac{1}{2}\log\frac{2\pi}{\eta_2} + \frac{\eta_1^2}{2\eta_2}.$$

Examples of Distributions in the Exponential Family

Gaussian:

$$\mathcal{N}(x|\mu,\sigma^2) = 1/\sqrt{2\pi\sigma^2} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$$

Exponential form:

$$\mathcal{N}(x|\mu, \sigma^2) = \exp\left(\boldsymbol{\eta}^\mathsf{T} \mathbf{u}(x) - g(\boldsymbol{\eta})\right)$$

$$\eta = (\mu/\sigma^2, 1.0/\sigma^2)^{\mathsf{T}}, \quad \mathbf{u}(x) = (x, -0.5x^2)^{\mathsf{T}}, \quad g(\eta) = -\frac{1}{2}\log\frac{2\pi}{\eta_2} + \frac{\eta_1^2}{2\eta_2}.$$

Most parametric distributions belong to the exponential family!

$$\begin{split} p_1(x) &= \frac{1}{\sqrt{2\pi\sigma_1^2}} \exp\left\{-\frac{1}{2}\sigma_1^2(x-\mu_1)^2\right\} \,, \\ p_2(x) &= \frac{1}{\sqrt{2\pi\sigma_2^2}} \exp\left\{-\frac{1}{2}\sigma_2^2(x-\mu_2)^2\right\} \,. \end{split}$$

Consider these two Gaussian distributions:

$$p_1(x) = \frac{1}{\sqrt{2\pi\sigma_1^2}} \exp\left\{-\frac{1}{2}\sigma_1^2(x-\mu_1)^2\right\},$$

$$p_2(x) = \frac{1}{\sqrt{2\pi\sigma_2^2}} \exp\left\{-\frac{1}{2}\sigma_2^2(x-\mu_2)^2\right\}.$$

• $p_1(x)p_2(x)$ is Gaussian with natural parameters $\eta_1 + \eta_2$.

$$p_1(x) = rac{1}{\sqrt{2\pi\sigma_1^2}} \exp\left\{-rac{1}{2}\sigma_1^2(x-\mu_1)^2
ight\} \,,$$
 $p_2(x) = rac{1}{\sqrt{2\pi\sigma_2^2}} \exp\left\{-rac{1}{2}\sigma_2^2(x-\mu_2)^2
ight\} \,.$

- $p_1(x)p_2(x)$ is Gaussian with natural parameters $\eta_1 + \eta_2$.
- The log-normalization constant of $p_1(x)p_2(x)$ is $g(\eta_1 + \eta_2) g(\eta_1) g(\eta_2)$.

$$p_1(x) = \frac{1}{\sqrt{2\pi\sigma_1^2}} \exp\left\{-\frac{1}{2}\sigma_1^2(x-\mu_1)^2\right\},$$

$$p_2(x) = \frac{1}{\sqrt{2\pi\sigma_2^2}} \exp\left\{-\frac{1}{2}\sigma_2^2(x-\mu_2)^2\right\}.$$

- $p_1(x)p_2(x)$ is Gaussian with natural parameters $\eta_1 + \eta_2$.
- The log-normalization constant of $p_1(x)p_2(x)$ is $g(\eta_1 + \eta_2) g(\eta_1) g(\eta_2)$.
- $p_1(x)/p_2(x)$ is Gaussian with natural parameters $\eta_1-\eta_2$.

$$p_1(x) = rac{1}{\sqrt{2\pi\sigma_1^2}} \exp\left\{-rac{1}{2}\sigma_1^2(x-\mu_1)^2\right\} ,$$
 $p_2(x) = rac{1}{\sqrt{2\pi\sigma_2^2}} \exp\left\{-rac{1}{2}\sigma_2^2(x-\mu_2)^2\right\} .$

- $p_1(x)p_2(x)$ is Gaussian with natural parameters $\eta_1 + \eta_2$.
- The log-normalization constant of $p_1(x)p_2(x)$ is $g(\eta_1 + \eta_2) g(\eta_1) g(\eta_2)$.
- $p_1(x)/p_2(x)$ is Gaussian with natural parameters $\eta_1-\eta_2$.
- The log-normalization constant of $p_1(x)/p_2(x)$ is $g(\eta_1 \eta_2) g(\eta_1) + g(\eta_2)$.

Consider the KL-divergence between p and q (q in the exponential family):

$$\mathsf{KL}(p||q) = -\int p(\mathbf{x}) \log \left\{ rac{q(\mathbf{x})}{p(\mathbf{x})}
ight\} d\mathbf{x} = g(oldsymbol{\eta}) - oldsymbol{\eta}^\mathsf{T} \mathbb{E}_p[\mathbf{u}(\mathbf{x})] + \mathsf{Const} \,.$$

Consider the KL-divergence between p and q (q in the exponential family):

$$\mathsf{KL}(p||q) = -\int p(\mathbf{x}) \log \left\{ rac{q(\mathbf{x})}{p(\mathbf{x})}
ight\} d\mathbf{x} = g(oldsymbol{\eta}) - oldsymbol{\eta}^\mathsf{T} \mathbb{E}_p[\mathbf{u}(\mathbf{x})] + \mathsf{Const} \,.$$

When **minimizing** with respect to the natural parameters η of q:

$$\frac{\partial \mathsf{KL}(p||q)}{\partial \boldsymbol{\eta}} = 0 \Longleftrightarrow \frac{\partial g(\boldsymbol{\eta})}{\partial \boldsymbol{\eta}} = \mathbb{E}_p[\mathbf{u}(\mathbf{x})],$$

Consider the KL-divergence between p and q (q in the exponential family):

$$\mathsf{KL}(p||q) = -\int p(\mathbf{x}) \log \left\{ rac{q(\mathbf{x})}{p(\mathbf{x})}
ight\} d\mathbf{x} = g(oldsymbol{\eta}) - oldsymbol{\eta}^\mathsf{T} \mathbb{E}_p[\mathbf{u}(\mathbf{x})] + \mathsf{Const} \,.$$

When **minimizing** with respect to the natural parameters η of q:

$$\frac{\partial \mathsf{KL}(p||q)}{\partial \boldsymbol{\eta}} = 0 \Longleftrightarrow \frac{\partial g(\boldsymbol{\eta})}{\partial \boldsymbol{\eta}} = \mathbb{E}_p[\mathbf{u}(\mathbf{x})],$$

Furthermore, it is possible to show that:

$$rac{\partial g(oldsymbol{\eta})}{\partial oldsymbol{\eta}} = \mathbb{E}_q[\mathbf{u}(\mathbf{x})]\,.$$

Consider the KL-divergence between p and q (q in the exponential family):

$$\mathsf{KL}(p||q) = -\int p(\mathbf{x}) \log \left\{ rac{q(\mathbf{x})}{p(\mathbf{x})}
ight\} d\mathbf{x} = g(oldsymbol{\eta}) - oldsymbol{\eta}^\mathsf{T} \mathbb{E}_p[\mathbf{u}(\mathbf{x})] + \mathsf{Const} \,.$$

When **minimizing** with respect to the natural parameters η of q:

$$\frac{\partial \mathsf{KL}(p||q)}{\partial \boldsymbol{\eta}} = 0 \Longleftrightarrow \frac{\partial g(\boldsymbol{\eta})}{\partial \boldsymbol{\eta}} = \mathbb{E}_p[\mathbf{u}(\mathbf{x})],$$

Furthermore, it is possible to show that:

$$rac{\partial g(oldsymbol{\eta})}{\partial oldsymbol{\eta}} = \mathbb{E}_q[oldsymbol{\mathsf{u}}(oldsymbol{\mathsf{x}})]\,.$$

 $\mathsf{KL}(p||q)$ is minimized by matching expected sufficient statistics.

Consider the KL-divergence between p and q (q in the exponential family):

$$\mathsf{KL}(p||q) = -\int p(\mathbf{x}) \log \left\{ rac{q(\mathbf{x})}{p(\mathbf{x})}
ight\} d\mathbf{x} = g(oldsymbol{\eta}) - oldsymbol{\eta}^\mathsf{T} \mathbb{E}_p[\mathbf{u}(\mathbf{x})] + \mathsf{Const} \,.$$

When **minimizing** with respect to the natural parameters η of q:

$$\frac{\partial \mathsf{KL}(p||q)}{\partial \boldsymbol{\eta}} = 0 \Longleftrightarrow \frac{\partial g(\boldsymbol{\eta})}{\partial \boldsymbol{\eta}} = \mathbb{E}_p[\mathbf{u}(\mathbf{x})],$$

Furthermore, it is possible to show that:

$$rac{\partial g(oldsymbol{\eta})}{\partial oldsymbol{\eta}} = \mathbb{E}_q[oldsymbol{\mathsf{u}}(oldsymbol{\mathsf{x}})]\,.$$

$\mathsf{KL}(p||q)$ is minimized by matching expected sufficient statistics.

If
$$q$$
 is Gaussian, then we have to match $\mathbb{E}_q[\mathbf{x}] = \mathbb{E}_p[\mathbf{x}]$ and $\mathbb{E}_q[\mathbf{x}\mathbf{x}^\mathsf{T}] = \mathbb{E}_p[\mathbf{x}\mathbf{x}^\mathsf{T}].$

Joint Approximation

EP approximates this joint distribution by a product of simpler factors:

$$p(\mathbf{f},\mathbf{y}) = \prod_{i=1}^{N} \phi_i(y_i f(\mathbf{x}_i) \mathcal{N}(\mathbf{f}|\mathbf{0}, \tilde{\mathbf{\Sigma}})) = \prod_i t_i(\mathbf{f}) \approx \prod_i \tilde{t}_i(\mathbf{f}),$$

where each \tilde{t}_i approximates the corresponding t_i . Each \tilde{t}_i must belong to the exponential family but need not be normalized.

Joint Approximation

EP approximates this joint distribution by a product of simpler factors:

$$p(\mathbf{f}, \mathbf{y}) = \prod_{i=1}^{N} \phi_i(y_i f(\mathbf{x}_i) \mathcal{N}(\mathbf{f} | \mathbf{0}, \tilde{\Sigma})) = \prod_i t_i(\mathbf{f}) \approx \prod_i \tilde{t}_i(\mathbf{f}),$$

where each \tilde{t}_i approximates the corresponding t_i . Each \tilde{t}_i must belong to the exponential family but need not be normalized.

The exponential family is closed under the product and $\prod_i \tilde{t}_i$ can be easily normalized to compute an approximate distribution:

$$p(\mathbf{f}|\mathbf{y}) = \frac{1}{p(\mathbf{y})} \prod_i t_i(\mathbf{f}) \approx \frac{1}{Z} \prod_i \tilde{t}_i(\mathbf{f}) = q(\mathbf{f}),$$

where $Z = \int \prod_i \tilde{t}_i(\mathbf{f}) d\mathbf{f}$ can be used to **approximate** $p(\mathbf{y})$.

Joint Approximation

EP approximates this joint distribution by a product of simpler factors:

$$p(\mathbf{f}, \mathbf{y}) = \prod_{i=1}^{N} \phi_i(y_i f(\mathbf{x}_i) \mathcal{N}(\mathbf{f}|\mathbf{0}, \tilde{\Sigma})) = \prod_i t_i(\mathbf{f}) \approx \prod_i \tilde{t}_i(\mathbf{f}),$$

where each \tilde{t}_i approximates the corresponding t_i . Each \tilde{t}_i must belong to the exponential family but need not be normalized.

The exponential family is closed under the product and $\prod_i \tilde{t}_i$ can be easily normalized to compute an approximate distribution:

$$p(\mathbf{f}|\mathbf{y}) = \frac{1}{p(\mathbf{y})} \prod_i t_i(\mathbf{f}) \approx \frac{1}{Z} \prod_i \tilde{t}_i(\mathbf{f}) = q(\mathbf{f}),$$

where $Z = \int \prod_i \tilde{t}_i(\mathbf{f}) d\mathbf{f}$ can be used to **approximate** $p(\mathbf{y})$.

Therefore q has the same form as the approximate factors!

How do we determine each approximate factor \tilde{t}_i ?

How do we determine each approximate factor \tilde{t}_i ?

We would like to **minimize** KL(p||q), but this is **intractable**!

How do we determine each approximate factor \tilde{t}_i ?

We would like to **minimize** KL(p||q), but this is **intractable**!

EP minimizes the KL divergence **between pairs** of t_i and \tilde{t}_i . This has the risk that **the product** may not be a good approximation. EP tries to **circumvent** this by an iterative procedure.

How do we determine each approximate factor \tilde{t}_i ?

We would like to **minimize** KL(p||q), but this is **intractable**!

EP minimizes the KL divergence **between pairs** of t_i and \tilde{t}_i . This has the risk that **the product** may not be a good approximation. EP tries to **circumvent** this by an iterative procedure.

Suppose we wish to refine \tilde{t}_j . We first remove this factor from the product:

$$q^{\bigvee j}(\mathbf{f}) \propto \prod_{i \neq j} \tilde{t}_i(\mathbf{f}) \propto q(\mathbf{f})/\tilde{t}_j(\mathbf{f})\,,$$

How do we determine each approximate factor \tilde{t}_i ?

We would like to **minimize** KL(p||q), but this is **intractable**!

EP minimizes the KL divergence **between pairs** of t_i and \tilde{t}_i . This has the risk that **the product** may not be a good approximation. EP tries to **circumvent** this by an iterative procedure.

Suppose we wish to refine \tilde{t}_j . We first remove this factor from the product:

$$q^{\setminus j}(\mathbf{f}) \propto \prod_{i \neq j} \tilde{t}_i(\mathbf{f}) \propto q(\mathbf{f})/\tilde{t}_j(\mathbf{f}),$$

Then, \tilde{t}_j is updated to minimize the KL-divergence between:

$$q_{\mathsf{new}}(\mathbf{f}) \propto \widetilde{t}_j(\mathbf{f}) q^{\setminus j}(\mathbf{f}) \,, \quad \hat{p}_j(\mathbf{f}) = rac{1}{Z_i} t_j(\mathbf{f}) q^{\setminus j}(\mathbf{f}) \,, \quad Z_j = \int t_j(\mathbf{f}) q^{\setminus j}(\mathbf{f}) d\mathbf{f} \,,$$

where $q^{\setminus j}$ is fixed. This ensures that \tilde{t}_i is accurate where $q^{\setminus j}$ is high.

In practice, \tilde{t}_j is found by first **minimizing** with respect to q_{new} :

$$\mathsf{KL}\left(\left. rac{t_j(\mathbf{f})q^{\setminus j}(\mathbf{f})}{Z_j} \right| q_{\mathsf{new}}(\mathbf{f})
ight) \, .$$

In practice, \tilde{t}_j is found by first **minimizing** with respect to q_{new} :

$$\mathsf{KL}\left(\left. rac{t_j(\mathbf{f})q^{\setminus j}(\mathbf{f})}{Z_j} \right| q_{\mathsf{new}}(\mathbf{f})
ight) \, .$$

This is done by matching expected sufficient statistics. As q is Gaussian, we only have to match the mean and the variance.

In practice, \tilde{t}_j is found by first **minimizing** with respect to q_{new} :

$$\mathsf{KL}\left(\left. rac{t_j(\mathbf{f})q^{\setminus j}(\mathbf{f})}{Z_j} \right| q_\mathsf{new}(\mathbf{f})
ight) \,.$$

This is done by **matching expected sufficient statistics**. As q is Gaussian, we only have to match the mean and the variance.

It is required that the moments of $\hat{p}_j(\mathbf{f}) = 1/Z_j f_j(\mathbf{f}) q^{\setminus j}(\mathbf{f})$ are tractable.

In practice, \tilde{t}_j is found by first **minimizing** with respect to q_{new} :

$$\mathsf{KL}\left(\left. rac{t_j(\mathbf{f})q^{\setminus j}(\mathbf{f})}{Z_j} \right| q_\mathsf{new}(\mathbf{f})
ight) \,.$$

This is done by **matching expected sufficient statistics**. As q is Gaussian, we only have to match the mean and the variance.

It is required that the moments of $\hat{p}_j(\mathbf{f}) = 1/Z_j f_j(\mathbf{f}) q^{\setminus j}(\mathbf{f})$ are tractable.

The refined factor \tilde{t}_j is set in practice to be:

$$ilde{t}_j(\mathbf{f}) = Z_j rac{q_{\mathsf{new}}(\mathbf{f})}{q^{\setminus j}(\mathbf{f})} \,, \qquad \qquad \mathsf{with} \quad ilde{t}_j(\mathbf{f}) q^{\setminus j}(\mathbf{f}) \propto q_{\mathsf{new}} \,,$$

which ensures that $\tilde{t}_j(\mathbf{f})q^{\vee j}(\mathbf{f})$ and $t_j(\mathbf{f})q^{\vee j}(\mathbf{f})$ integrate the same.

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of q.

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of q.

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of q.

EP Algorithm in General: Computes q and an approximation to $p(\mathbf{y})$.

1 Initialize q and each \tilde{t}_i to be uniform.

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of q.

- 1 Initialize q and each \tilde{t}_i to be uniform.
- **2** Repeat until convergence of the \tilde{t}_i :

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of q.

- 1 Initialize q and each \tilde{t}_i to be uniform.
- **2** Repeat until convergence of the \tilde{t}_i :
 - **1** Choose a factor \tilde{t}_j to refine.

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of q.

- 1 Initialize q and each \tilde{t}_i to be uniform.
- **2** Repeat until convergence of the \tilde{t}_i :
 - **1** Choose a factor \tilde{t}_i to refine.
 - **2** Remove \tilde{t}_j from \tilde{q} by division $q^{\setminus j} \propto q/\tilde{t}_j$.

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of q.

- 1 Initialize q and each \tilde{t}_i to be uniform.
- **2** Repeat until convergence of the \tilde{t}_i :
 - **1** Choose a factor \tilde{t}_i to refine.
 - **2** Remove \tilde{t}_j from \tilde{q} by division $q^{\setminus j} \propto q/\tilde{t}_j$.
 - **3** Compute Z_i and \hat{p}_i and find q_{new} by minimizing $\text{KL}(\hat{p}_i||q_{\text{new}})$.

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of q.

- 1 Initialize q and each \tilde{t}_i to be uniform.
- **2** Repeat until convergence of the \tilde{t}_i :
 - 1 Choose a factor \tilde{t}_i to refine.
 - **2** Remove \tilde{t}_j from \tilde{q} by division $q^{\setminus j} \propto q/\tilde{t}_j$.
 - **3** Compute Z_j and \hat{p}_j and find q_{new} by minimizing $\text{KL}(\hat{p}_j||q_{\text{new}})$.
 - **4** Compute and store the new factor $\tilde{t}_j = Z_j q_{\text{new}}/q^{\setminus j}$.

Full Algorithm

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of q.

EP Algorithm in General: Computes q and an approximation to $p(\mathbf{y})$.

- 1 Initialize q and each \tilde{t}_i to be uniform.
- **2** Repeat until convergence of the \tilde{t}_i :
 - 1 Choose a factor \tilde{t}_i to refine.
 - **2** Remove \tilde{t}_j from \tilde{q} by division $q^{\setminus j} \propto q/\tilde{t}_j$.
 - **3** Compute Z_j and \hat{p}_j and find q_{new} by minimizing $\text{KL}(\hat{p}_j||q_{\text{new}})$.
 - 4 Compute and store the new factor $\tilde{t}_i = Z_i q_{\text{new}}/q^{ij}$.
- **3** Evaluate the approximation to the model evidence:

$$p(\mathbf{y}) \approx Z = \int \prod_{j} \tilde{t}_{j}(\mathbf{f}) d\mathbf{f}$$
.

Full Algorithm

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of q.

EP Algorithm in General: Computes q and an approximation to $p(\mathbf{y})$.

- 1 Initialize q and each \tilde{t}_i to be uniform.
- **2** Repeat until convergence of the \tilde{t}_i :
 - **1** Choose a factor \tilde{t}_i to refine.
 - **2** Remove \tilde{t}_j from \tilde{q} by division $q^{\setminus j} \propto q/\tilde{t}_j$.
 - **3** Compute Z_j and \hat{p}_j and find q_{new} by minimizing $\text{KL}(\hat{p}_j||q_{\text{new}})$.
 - 4 Compute and store the new factor $\tilde{t}_i = Z_i q_{\text{new}}/q^{ij}$.
- 3 Evaluate the approximation to the model evidence:

$$p(\mathbf{y}) \approx Z = \int \prod_{j} \tilde{t}_{j}(\mathbf{f}) d\mathbf{f}$$
.

The FITC prior results in a total cost of $\mathcal{O}(NM^2)$!

Approximates
$$p(\mathbf{f}|\mathbf{y}) \propto t_0(\mathbf{f}) \prod_{j=1}^N t_j(\mathbf{f})$$
 with $q(\mathbf{f}) \propto t_0(\mathbf{f}) \prod_{j=1}^N \tilde{t}_j(\mathbf{t})$

Approximates
$$p(\mathbf{f}|\mathbf{y}) \propto t_0(\mathbf{f}) \prod_{j=1}^N t_j(\mathbf{f})$$
 with $q(\mathbf{f}) \propto t_0(\mathbf{f}) \prod_{j=1}^N \tilde{t}_j(\mathbf{t})$

$$p(\mathbf{f}|\mathbf{y}) \propto t_0(\mathbf{f}) \quad t_1(\mathbf{f}) \ t_2(\mathbf{f}) \ t_3(\mathbf{f}) \qquad \approx q(\mathbf{f}) \propto t_0(\mathbf{f}) \quad \tilde{t}_1(\mathbf{f}) \ \tilde{t}_2(\mathbf{f}) \ \tilde{t}_3(\mathbf{f})$$

Approximates
$$p(\mathbf{f}|\mathbf{y}) \propto t_0(\mathbf{f}) \prod_{j=1}^N t_j(\mathbf{f})$$
 with $q(\mathbf{f}) \propto t_0(\mathbf{f}) \prod_{j=1}^N \tilde{t}_j(\mathbf{t})$

$$p(\mathbf{f}|\mathbf{y}) \propto t_0(\mathbf{f}) \quad t_1(\mathbf{f}) \ t_2(\mathbf{f}) \ t_3(\mathbf{f}) \qquad \approx \qquad q(\mathbf{f}) \propto t_0(\mathbf{f}) \quad \tilde{t}_1(\mathbf{f}) \ \tilde{t}_2(\mathbf{f}) \ \tilde{t}_3(\mathbf{f})$$

The \tilde{t}_i are tuned by minimizing the KL-divergence

$$\mathsf{KL}[\hat{
ho}_j||q] \quad \mathsf{for}\, j=1,\ldots, \mathsf{N}\,, \quad \mathsf{where} \quad egin{aligned} \hat{
ho}_j(\mathbf{f}) & \propto & t_j(\mathbf{f}) \prod_{i
eq j} \, ilde{t}_i(\mathbf{f}) \ q(\mathbf{f}) & \propto & ilde{t}_j(\mathbf{f}) \prod_{i
eq j} \, ilde{t}_i(\mathbf{f}) \,. \end{aligned}$$

Approximates
$$p(\mathbf{f}|\mathbf{y}) \propto t_0(\mathbf{f}) \prod_{j=1}^N t_j(\mathbf{f})$$
 with $q(\mathbf{f}) \propto t_0(\mathbf{f}) \prod_{j=1}^N \tilde{t}_j(\mathbf{t})$

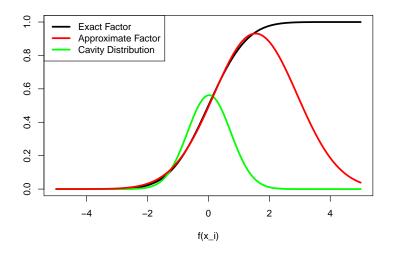
$$p(\mathbf{f}|\mathbf{y}) \propto t_0(\mathbf{f}) \quad t_1(\mathbf{f}) \ t_2(\mathbf{f}) \ t_3(\mathbf{f}) \qquad \approx \qquad q(\mathbf{f}) \propto t_0(\mathbf{f}) \quad \tilde{t}_1(\mathbf{f}) \ \tilde{t}_2(\mathbf{f}) \ \tilde{t}_3(\mathbf{f})$$

The \tilde{t}_i are tuned by minimizing the KL-divergence

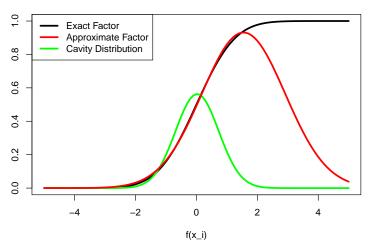
$$\mathsf{KL}[\hat{
ho}_j||q] \quad \mathsf{for}\, j=1,\ldots, \mathsf{N}\,, \quad \mathsf{where} \quad egin{aligned} \hat{
ho}_j(\mathbf{f}) & \propto & t_j(\mathbf{f}) \prod_{i
eq j} \, ilde{t}_i(\mathbf{f}) \ q(\mathbf{f}) & \propto & ilde{t}_j(\mathbf{f}) \prod_{i
eq j} \, ilde{t}_i(\mathbf{f}) \,. \end{aligned}$$

If the exact factor already belongs to the exponential family it needs not be approximated!

GFITC: Factor Approximation



GFITC: Factor Approximation



The approximate factor is accurate in regions of high posterior probability as indicated by the cavity distribution!

They are optimized by maximizing the EP estimate of the log-marginal likelihood $\log Z \approx \log p(y)$.

They are optimized by maximizing the EP estimate of the log-marginal likelihood $\log Z \approx \log p(y)$.

Problem: The parameters θ_i of the approximate factors also depend on the hyper-parameters (including the inducing points)!

They are optimized by maximizing the EP estimate of the log-marginal likelihood $\log Z \approx \log p(y)$.

Problem: The parameters θ_i of the approximate factors also depend on the hyper-parameters (including the inducing points)!

• Direct dependence of log Z on the hyper-parameters.

They are optimized by maximizing the EP estimate of the log-marginal likelihood $\log Z \approx \log p(y)$.

Problem: The parameters θ_i of the approximate factors also depend on the hyper-parameters (including the inducing points)!

- Direct dependence of log Z on the hyper-parameters.
- Indirect dependence of log Z on the hyper-parameters via each θ_i .

They are optimized by maximizing the EP estimate of the log-marginal likelihood $\log Z \approx \log p(y)$.

Problem: The parameters θ_i of the approximate factors also depend on the hyper-parameters (including the inducing points)!

- Direct dependence of log Z on the hyper-parameters.
- Indirect dependence of log Z on the hyper-parameters via each θ_i .

If EP converges the gradient of $\log Z$ w.r.t. each θ_i is zero, which allows to easily compute the gradients of $\log Z!$

We want to compute the value of f^* at a new x^* :

We want to compute the value of f^* at a new x^* :

$$p(\mathbf{f}^{\star}, \mathbf{f}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f}^{\star} \\ \mathbf{f} \end{array}\right] \middle| \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} & \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}} \\ \mathbf{Q}_{\mathbf{f}\mathbf{f}^{\star}} & \tilde{\mathbf{\Sigma}}_{\mathbf{f}\mathbf{f}} \end{array}\right]\right)$$

We want to compute the value of f^* at a new x^* :

$$p(\mathbf{f}^{\star}, \mathbf{f}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f}^{\star} \\ \mathbf{f} \end{array}\right] \middle| \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{c} \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} & \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}} \\ \mathbf{Q}_{\mathbf{f}\mathbf{f}^{\star}} & \tilde{\mathbf{\Sigma}}_{\mathbf{f}\mathbf{f}} \end{array}\right]\right)$$

The conditional $p(\mathbf{f}^*|\mathbf{f})$ is:

$$\textit{p}(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}|\ \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{f},\boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}^{\mathbf{f}}}^{-1}\mathbf{Q}_{\mathbf{f}^{\star}\star}\right)$$

We want to compute the value of f^* at a new x^* :

$$p(\mathbf{f}^{\star},\mathbf{f}) = \mathcal{N}\left(\left[\begin{array}{c}\mathbf{f}^{\star}\\\mathbf{f}\end{array}\right] \middle| \left[\begin{array}{c}\mathbf{0}\\\mathbf{0}\end{array}\right], \left[\begin{array}{cc}\boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} & \boldsymbol{Q}_{\mathbf{f}^{\star}\mathbf{f}}\\ \boldsymbol{Q}_{\mathbf{f}\mathbf{f}^{\star}} & \boldsymbol{\tilde{\Sigma}}_{\mathbf{f}\mathbf{f}}\end{array}\right]\right)$$

The conditional $p(\mathbf{f}^*|\mathbf{f})$ is:

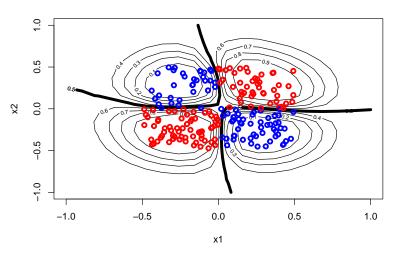
$$p(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}|\ \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{ff}}^{-1}\mathbf{f},\boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathbf{ff}}^{-1}\mathbf{Q}_{\mathbf{f}^{\star}}\right)$$

After marginalizing f w.r.t. q(f), we obtain the predictive distribution:

$$\begin{split} \rho(\mathbf{f}^{\star}|\mathbf{y}) &= \int \rho(\mathbf{f}^{\star}|\mathbf{f})q(\mathbf{f})d\mathbf{f} \\ &= \mathcal{N}\left(\mathbf{f}^{\star}|\ \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\left(\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}\mathbf{f}} + \tilde{\boldsymbol{\Pi}}\right)^{-1}\tilde{\mathbf{y}}, \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\left(\tilde{\boldsymbol{\Sigma}}_{\mathbf{f}\mathbf{f}} + \tilde{\boldsymbol{\Pi}}\right)^{-1}\mathbf{Q}_{\mathbf{f}\mathbf{f}^{\star}}\right) \end{split}$$

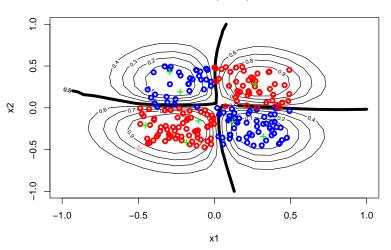
GFITC: Illustrative Example





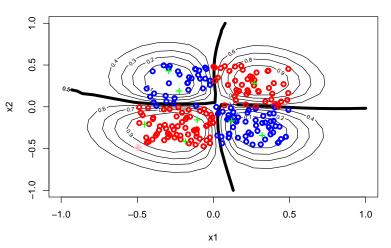
GFITC: Illustrative Example





GFITC: Illustrative Example





The inducing points spread across the input space!

Previous methods approximate the GP prior using a low rank approximation of Σ , resulting in a cost $\mathcal{O}(NM^2)$.

Previous methods approximate the GP prior using a low rank approximation of Σ , resulting in a cost $\mathcal{O}(NM^2)$.

Previous methods approximate the GP prior using a low rank approximation of Σ , resulting in a cost $\mathcal{O}(NM^2)$.

Variational Free Energy (VFE) Method:

• Keeps the GP prior intact and does not introduce any simplification!

Previous methods approximate the GP prior using a low rank approximation of Σ , resulting in a cost $\mathcal{O}(NM^2)$.

- Keeps the GP prior intact and does not introduce any simplification!
- Carries out approximate inference to approximate the GP posterior.

Previous methods approximate the GP prior using a low rank approximation of Σ , resulting in a cost $\mathcal{O}(NM^2)$.

- Keeps the GP prior intact and does not introduce any simplification!
- Carries out approximate inference to approximate the GP posterior.
- The particular approximate distribution q results in cost $\mathcal{O}(NM^2)$.

Previous methods approximate the GP prior using a low rank approximation of Σ , resulting in a cost $\mathcal{O}(NM^2)$.

- Keeps the GP prior intact and does not introduce any simplification!
- Carries out approximate inference to approximate the GP posterior.
- The particular approximate distribution q results in cost $\mathcal{O}(NM^2)$.
- Variational inference is used to tune q.

Previous methods approximate the GP prior using a low rank approximation of Σ , resulting in a cost $\mathcal{O}(NM^2)$.

Variational Free Energy (VFE) Method:

- Keeps the GP prior intact and does not introduce any simplification!
- Carries out approximate inference to approximate the GP posterior.
- The particular approximate distribution q results in cost $\mathcal{O}(NM^2)$.
- Variational inference is used to tune q.

Since the GP prior is not changed it tends to perform better than the previous methods!

Adjust the parameters of q to match p by minimizing $KL(q|p) \ge 0$.

Adjust the parameters of q to match p by minimizing $KL(q|p) \ge 0$.

$$\mathsf{KL}(q|p) = 0 \Longleftrightarrow q(\mathbf{f}) = p(\mathbf{f})$$

Adjust the parameters of q to match p by minimizing $KL(q|p) \ge 0$.

$$\mathsf{KL}(q|p) = 0 \Longleftrightarrow q(\mathbf{f}) = p(\mathbf{f})$$

The expression for the KL divergence between q and p is:

$$\int q(\mathbf{f})\log\frac{q(\mathbf{f})}{p(\mathbf{f})}d\mathbf{f}\geq 0$$

Adjust the parameters of q to match p by minimizing $KL(q|p) \ge 0$.

$$\mathsf{KL}(q|p) = 0 \Longleftrightarrow q(\mathbf{f}) = p(\mathbf{f})$$

The expression for the KL divergence between q and p is:

$$\int q(\mathbf{f})\log\frac{q(\mathbf{f})}{p(\mathbf{f})}d\mathbf{f}\geq 0$$

KL(q|p) depends on p, which is assumed to be intractable!

Adjust the parameters of q to match p by minimizing $KL(q|p) \ge 0$.

$$\mathsf{KL}(q|p) = 0 \Longleftrightarrow q(\mathbf{f}) = p(\mathbf{f})$$

The expression for the KL divergence between q and p is:

$$\int q(\mathbf{f})\log\frac{q(\mathbf{f})}{p(\mathbf{f})}d\mathbf{f}\geq 0$$

 $\mathsf{KL}(q|p)$ depends on p, which is assumed to be intractable!

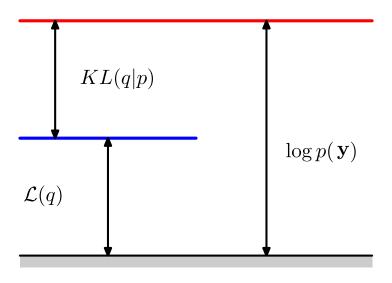
Let the target be $p(\mathbf{f}|\mathbf{y})$. Consider the decomposition of $p(\mathbf{y})$:

$$\log p(\mathbf{y}) = \mathcal{L}(q) + \mathsf{KL}(q|p),$$

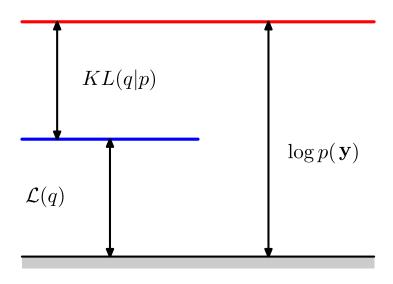
where

$$\mathcal{L}(q) = \int q(\mathbf{f}) \log \frac{p(\mathbf{f}, \mathbf{y})}{q(\mathbf{f})} d\mathbf{f}, \quad \mathsf{KL}(q|p) = \int q(\mathbf{f}) \log \frac{q(\mathbf{f})}{p(\mathbf{f}|\mathbf{y})} d\mathbf{f}.$$

Decomposition of the Marginal Likelihood



Decomposition of the Marginal Likelihood

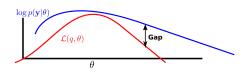


 $\mathcal{L}(q)$ can be used to approximate $\log p(\mathbf{y})$ if $\mathrm{KL}(q|p)$ is small!

Variational Free Energy (VFE)

Lower bound the log-likelihood:

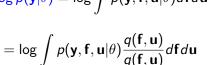
$$\log p(\mathbf{y}|\theta) = \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta) d\mathbf{f} d\mathbf{u}$$

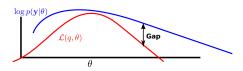


Variational Free Energy (VFE)

Lower bound the log-likelihood:

$$\log p(\mathbf{y}|\theta) = \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta) d\mathbf{f} d\mathbf{u}$$





Lower bound the log-likelihood:

$$\log p(\mathbf{y}|\theta) = \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta) d\mathbf{f} d\mathbf{u}$$

$$\log p(\mathbf{y}|\theta)$$
 $\mathcal{L}(q,\theta)$ Gap

$$= \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta) \frac{q(\mathbf{f}, \mathbf{u})}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u} \ge \int q(\mathbf{f}, \mathbf{u}) \log \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta)}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u} \equiv \mathcal{L}(\mathbf{q}, \theta)$$

Lower bound the log-likelihood:

$$\log p(\mathbf{y}|\theta) = \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta) d\mathbf{f} d\mathbf{u}$$

$$\log p(\mathbf{y}|\theta)$$

$$\mathcal{L}(q,\theta)$$
 Gap

$$= \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u} | \theta) \frac{q(\mathbf{f}, \mathbf{u})}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u} \ge \int q(\mathbf{f}, \mathbf{u}) \log \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u} | \theta)}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u} \equiv \mathcal{L}(\mathbf{q}, \theta)$$

$$\mathcal{L}(q,\theta) = \int q(\mathbf{f},\mathbf{u}) \log \frac{p(\mathbf{y},\mathbf{f},\mathbf{u}|\theta)}{q(\mathbf{f},\mathbf{u})} d\mathbf{f} d\mathbf{u} = \log p(\mathbf{y}|\theta) - \mathsf{KL}[q(\mathbf{f},\mathbf{u})|p(\mathbf{f},\mathbf{u}|\mathbf{y})]$$

Lower bound the log-likelihood:

$$\log p(\mathbf{y}|\theta) = \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta) d\mathbf{f} d\mathbf{u}$$

$$\frac{\log p(\mathbf{y}|\theta)}{\mathcal{L}(q,\theta)}$$
 Gap

$$= \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u} | \theta) \frac{q(\mathbf{f}, \mathbf{u})}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u} \ge \int q(\mathbf{f}, \mathbf{u}) \log \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u} | \theta)}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u} \equiv \mathcal{L}(\mathbf{q}, \theta)$$

$$\mathcal{L}(q,\theta) = \int q(\mathbf{f},\mathbf{u}) \log \frac{p(\mathbf{y},\mathbf{f},\mathbf{u}|\theta)}{q(\mathbf{f},\mathbf{u})} d\mathbf{f} d\mathbf{u} = \log p(\mathbf{y}|\theta) - \mathsf{KL}[q(\mathbf{f},\mathbf{u})|p(\mathbf{f},\mathbf{u}|\mathbf{y})]$$

KL ≡ Kullback-Leibler divergence

Lower bound the log-likelihood:

$$\log p(\mathbf{y}|\theta) = \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta) d\mathbf{f} d\mathbf{u}$$

$$\frac{\log p(\mathbf{y}|\theta)}{\mathcal{L}(q,\theta)}$$

$$= \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u} | \theta) \frac{q(\mathbf{f}, \mathbf{u})}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u} \ge \int q(\mathbf{f}, \mathbf{u}) \log \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u} | \theta)}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u} \equiv \frac{\mathcal{L}(\mathbf{q}, \theta)}{q(\mathbf{f}, \mathbf{u})}$$

$$\mathcal{L}(q,\theta) = \int q(\mathbf{f},\mathbf{u}) \log \frac{p(\mathbf{y},\mathbf{f},\mathbf{u}|\theta)}{q(\mathbf{f},\mathbf{u})} d\mathbf{f} d\mathbf{u} = \log p(\mathbf{y}|\theta) - \mathsf{KL}[q(\mathbf{f},\mathbf{u})|p(\mathbf{f},\mathbf{u}|\mathbf{y})]$$

KL ≡ Kullback-Leibler divergence

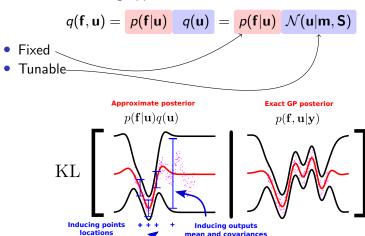
By maximizing $\mathcal{L}(q, \theta)$ w.r.t q we are enforcing that $q(f, \mathbf{u})$ looks similar to $p(f, \mathbf{u}|\mathbf{y})$ in terms of the KL!

Consider the following approximate distribution:

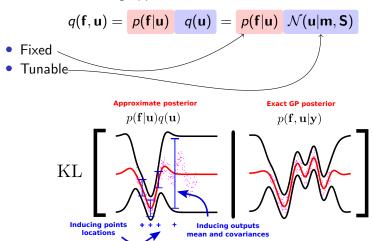
Consider the following approximate distribution:

$$q(\mathbf{f}, \mathbf{u}) = \boxed{p(\mathbf{f}|\mathbf{u})} \quad q(\mathbf{u}) = \boxed{p(\mathbf{f}|\mathbf{u})} \quad \mathcal{N}(\mathbf{u}|\mathbf{m}, \mathbf{S})$$
• Fixed
• Tunable

Consider the following approximate distribution:



Consider the following approximate distribution:



The inducing points are now parameters of the approx. dist. q!

$$\mathcal{L}(q,\theta) = \int q(\mathbf{f}, \mathbf{u}) \log \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta)}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u}$$
$$= \int p(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \frac{p(\mathbf{y}|\mathbf{f}, \theta) p(\mathbf{f}|\mathbf{u}) p(\mathbf{u})}{p(\mathbf{f}|\mathbf{u}) q(\mathbf{u})} d\mathbf{f} d\mathbf{u}$$

$$\mathcal{L}(q,\theta) = \int q(\mathbf{f},\mathbf{u}) \log \frac{p(\mathbf{y},\mathbf{f},\mathbf{u}|\theta)}{q(\mathbf{f},\mathbf{u})} d\mathbf{f} d\mathbf{u}$$

$$= \int p(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \frac{p(\mathbf{y}|\mathbf{f},\theta) p(\mathbf{f}|\mathbf{u}) p(\mathbf{u})}{p(\mathbf{f}|\mathbf{u}) q(\mathbf{u})} d\mathbf{f} d\mathbf{u}$$

$$\mathcal{L}(q,\theta) = \int q(\mathbf{f}, \mathbf{u}) \log \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta)}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u}$$

$$= \int p(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \frac{p(\mathbf{y}|\mathbf{f}, \theta) p(\mathbf{f}|\mathbf{u}) p(\mathbf{u})}{p(\mathbf{f}|\mathbf{u}) q(\mathbf{u})} d\mathbf{f} d\mathbf{u}$$

$$\mathcal{L}(q,\theta) = \mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f},\theta)] - \mathbb{KL}[q(\mathbf{u})|p(\mathbf{u})]$$

- Mean squared prediction error
- KL between Gaussians -

$$\mathcal{L}(q,\theta) = \int q(\mathbf{f}, \mathbf{u}) \log \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta)}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u}$$

$$= \int p(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \frac{p(\mathbf{y}|\mathbf{f}, \theta) p(\mathbf{f}|\mathbf{u}) p(\mathbf{u})}{p(\mathbf{f}|\mathbf{u}) q(\mathbf{u})} d\mathbf{f} d\mathbf{u}$$

$$\mathcal{L}(q,\theta) = \mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f},\theta)] - \mathbb{KL}[q(\mathbf{u})|p(\mathbf{u})]$$

- Mean squared prediction error
- KL between Gaussians -
- No change in the model is made and the cost is in $\mathcal{O}(M^2N)$!

$$\mathcal{L}(q,\theta) = \int q(\mathbf{f}, \mathbf{u}) \log \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta)}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u}$$

$$= \int p(\mathbf{f}|\mathbf{u}) q(\mathbf{u}) \log \frac{p(\mathbf{y}|\mathbf{f}, \theta) p(\mathbf{f}|\mathbf{u}) p(\mathbf{u})}{p(\mathbf{f}|\mathbf{u}) q(\mathbf{u})} d\mathbf{f} d\mathbf{u}$$

$$\mathcal{L}(q,\theta) = \mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f},\theta)] - \mathbb{KL}[q(\mathbf{u})|p(\mathbf{u})]$$

- Mean squared prediction error
- KL between Gaussians -
- No change in the model is made and the cost is in $\mathcal{O}(M^2N)$!
- Predictions are made using $p(\mathbf{f}^*|\mathbf{u})q(\mathbf{u})$ marginalizing out \mathbf{u} .

We want to compute the value of f^* at a new x^* :

We want to compute the value of f^* at a new x^* :

$$p(\mathbf{f}^{\star}, \mathbf{u}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f}^{\star} \\ \mathbf{u} \end{array}\right] \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} & \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{u}} \\ \mathbf{\Sigma}_{\mathbf{u}\mathbf{f}^{\star}} & \mathbf{\Sigma}_{\mathbf{u}\mathbf{u}} \end{array}\right]\right)$$

We want to compute the value of f^* at a new x^* :

$$p(\mathbf{f}^{\star}, \mathbf{u}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f}^{\star} \\ \mathbf{u} \end{array}\right] \middle| \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array}\right], \left[\begin{array}{cc} \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} & \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{u}} \\ \mathbf{\Sigma}_{\mathbf{u}\mathbf{f}^{\star}} & \mathbf{\Sigma}_{\mathbf{u}\mathbf{u}} \end{array}\right]\right)$$

The conditional $p(\mathbf{f}^*|\mathbf{u})$ is:

$$\textit{p}(\textbf{f}^{\star}|\textbf{u}) = \mathcal{N}\left(\textbf{f}^{\star}|\ \boldsymbol{\Sigma}_{\textbf{f}^{\star}\textbf{u}}\boldsymbol{\Sigma}_{\textbf{u}\textbf{u}}^{-1}\textbf{u}, \boldsymbol{\Sigma}_{\textbf{f}^{\star}\textbf{f}^{\star}} - \boldsymbol{\Sigma}_{\textbf{f}^{\star}\textbf{u}}\boldsymbol{\Sigma}_{\textbf{u}\textbf{u}}^{-1}\boldsymbol{\Sigma}_{\textbf{u}\textbf{f}^{\star}}\right)$$

We want to compute the value of f^* at a new x^* :

$$p(\mathbf{f}^{\star},\mathbf{u}) = \mathcal{N}\left(\left[\begin{array}{c}\mathbf{f}^{\star}\\\mathbf{u}\end{array}\right] \middle| \left[\begin{array}{c}\mathbf{0}\\\mathbf{0}\end{array}\right], \left[\begin{array}{cc}\boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} & \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{u}}\\\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{f}^{\star}} & \boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}\end{array}\right]\right)$$

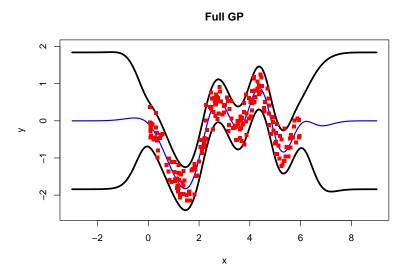
The conditional $p(\mathbf{f}^*|\mathbf{u})$ is:

$$\rho(\mathbf{f}^{\star}|\mathbf{u}) = \mathcal{N}\left(\mathbf{f}^{\star}|\ \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{u}}\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}^{-1}\mathbf{u}, \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{u}}\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}^{-1}\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{f}^{\star}}\right)$$

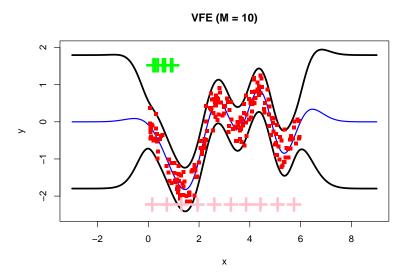
After marginalizing \mathbf{u} w.r.t. $q(\mathbf{u})$, we obtain the predictive distribution:

$$\begin{split} \rho(\mathbf{f}^{\star}|\mathbf{y}) &= \int \rho(\mathbf{f}^{\star}|\mathbf{u})q(\mathbf{u})d\mathbf{u} \\ &= \mathcal{N}\left(\mathbf{f}^{\star}|\ \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{u}}\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}^{-1}\mathbf{m},\boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{u}}\left(\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}^{-1} - \boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}^{-1}\mathbf{S}\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}^{-1}\right)\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{f}^{\star}}\right) \end{split}$$

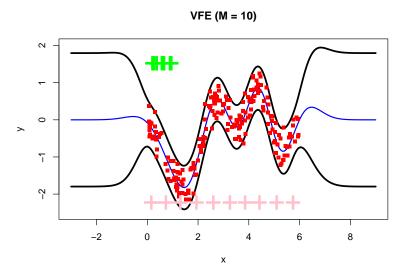
VFE: Illustrative Example



VFE: Illustrative Example



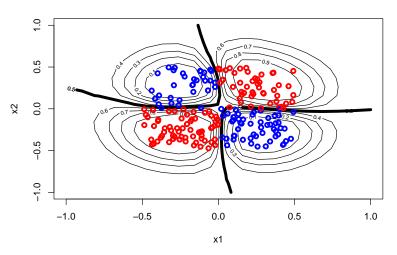
VFE: Illustrative Example



The inducing points cover the regions where the function changes!

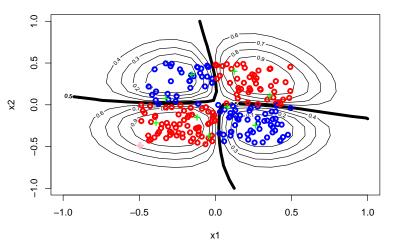
VFE: Illustrative Classification Example



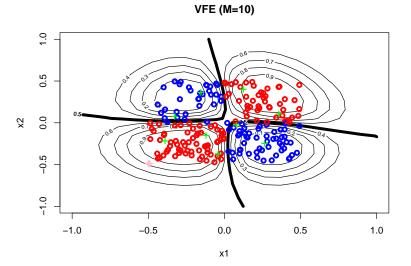


VFE: Illustrative Classification Example





VFE: Illustrative Classification Example



The inducing points spread across the input space!

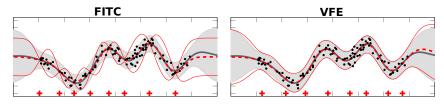
Two approaches:

Two approaches:

- FITC: optimize the marginal likelihood of an approximate GP model.
- VFE: maximize fidelity to the original exact GP.

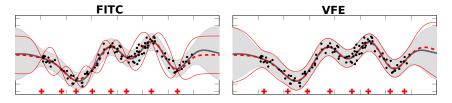
Two approaches:

- FITC: optimize the marginal likelihood of an approximate GP model.
- VFE: maximize fidelity to the original exact GP.



Two approaches:

- FITC: optimize the marginal likelihood of an approximate GP model.
- VFE: maximize fidelity to the original exact GP.



- FITC: less local optima and easier to optimize, also less accurate.
- VFE: more accurate, more local optima, more difficult to optimize.

(Bui et al., 2017) (Bauer et al., 2016)

Run the cells of the notebook to fit a sparse GP using the VFE approximation and complete task 6!

Whitened Parameterization for VFE

Alternative VFE objective expected to be easier to optimize!

Whitened Parameterization for VFE

Alternative VFE objective expected to be easier to optimize!

Instead of making inference about \mathbf{u} , the whitened VFE objective makes inference about:

e such that
$$\mathbf{u} = \mathbf{Le}$$
, $\mathbf{e} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$,

with u the latent process values at the inducing points and $\textbf{L}^{\mathsf{T}}\textbf{L} = \Sigma_{uu}.$

Whitened Parameterization for VFE

Alternative VFE objective expected to be easier to optimize!

Instead of making inference about \mathbf{u} , the whitened VFE objective makes inference about:

$$\mathbf{e}$$
 such that $\mathbf{u} = \mathbf{L}\mathbf{e}$, $\mathbf{e} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$,

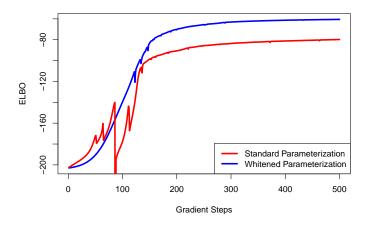
with u the latent process values at the inducing points and $\mathbf{L}^\mathsf{T}\mathbf{L} = \Sigma_{uu}$.

The VFE objective becomes:

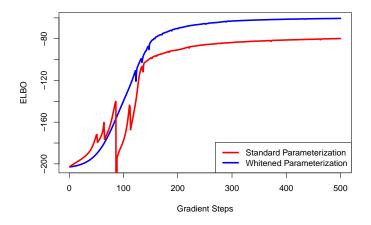
$$\sum_{i=1}^{N} \mathbb{E}_{q(\mathbf{e})p(f(\mathbf{x}_i)|\mathbf{e})} \left[\log p(y_i|f(\mathbf{x}_i)) \right] - \mathsf{KL}(q(\mathbf{e})|\mathcal{N}(\mathbf{0},\mathbf{I})),$$

with $p(f(\mathbf{x}_i)|\mathbf{e})$ using the covariances between $f(\mathbf{x}_i)$ and \mathbf{e} .

Whitened Parameterization: Illustrative Example



Whitened Parameterization: Illustrative Example



Whitening significantly improves convergence!

Run the cells of the notebook to fit a sparse GP using the VFE approximation with whitening!

Natural Gradient Ascent

Gradient ascent moves in the direction of the gradient $\nabla_{\xi} \mathcal{L}(\xi)$.

Gradient ascent moves in the direction of the gradient $\nabla_{\xi} \mathcal{L}(\xi)$.

Formally:

$$abla_{oldsymbol{\xi}} \mathcal{L}(oldsymbol{\xi}) \propto \lim_{\epsilon o 0} rac{1}{\epsilon} rg \max_{oldsymbol{d} \, s.t. ||oldsymbol{d}|| \leq \epsilon} \mathcal{L}(oldsymbol{\xi} + oldsymbol{d})$$

Gradient ascent moves in the direction of the gradient $\nabla_{\xi} \mathcal{L}(\xi)$.

Formally:

$$abla_{oldsymbol{\xi}} \mathcal{L}(oldsymbol{\xi}) \propto \lim_{\epsilon o 0} rac{1}{\epsilon} rg \max_{oldsymbol{d} \, s.t. ||oldsymbol{d}|| \leq \epsilon} \mathcal{L}(oldsymbol{\xi} + oldsymbol{d})$$

The steepest ascent direction picks d in the ϵ -vicinity of ξ (measured by the Euclidean norm) that maximizes $\mathcal{L}(\cdot)$.

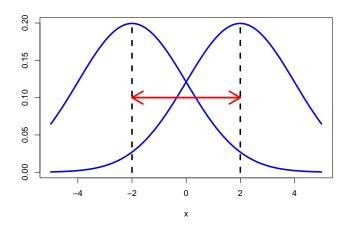
Gradient ascent moves in the direction of the gradient $\nabla_{\xi} \mathcal{L}(\xi)$.

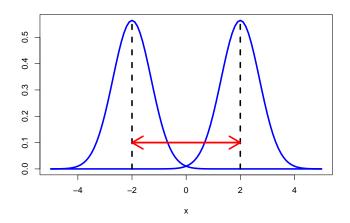
Formally:

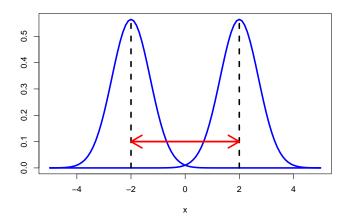
$$abla_{oldsymbol{\xi}} \mathcal{L}(oldsymbol{\xi}) \propto \lim_{\epsilon o 0} rac{1}{\epsilon} rg \max_{oldsymbol{d} \, s.t. || oldsymbol{d} || < \epsilon} \mathcal{L}(oldsymbol{\xi} + oldsymbol{d})$$

The steepest ascent direction picks d in the ϵ -vicinity of ξ (measured by the Euclidean norm) that maximizes $\mathcal{L}(\cdot)$.

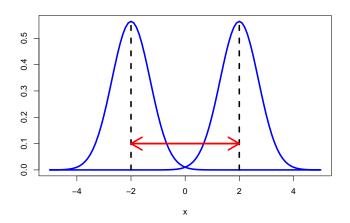
If ξ represents the parameters of probability distributions, the Euclidean norm may be problematic!







The Euclidean distance between parameters is 4 in both cases!



A better alternative is the KL-divergence between distributions!

Considers the KL-divergence as a norm:

$$\nabla_{\pmb{\xi}} \mathcal{L}(\pmb{\xi}) \mathbf{F}_{\pmb{\xi}}^{-1} \propto \lim_{\epsilon \to 0} \frac{1}{\epsilon} \underset{\mathbf{d} \, s.t. \mathsf{KL}[q(\mathbf{u}|\pmb{\xi})|q(\mathbf{u}|\pmb{\xi}+\mathbf{d})] \leq \epsilon}{\mathsf{arg max}} \mathcal{L}(\pmb{\xi} + \mathbf{d})$$

with $\mathbf{F}_{\mathcal{E}}$ the Fisher information of q:

$$\mathbf{F}_{\boldsymbol{\xi}} = -\mathbb{E}_{q(\mathbf{u}|\boldsymbol{\xi})}[\nabla_{\boldsymbol{\xi}}^2 \log q(\mathbf{u}|\boldsymbol{\xi})]$$

Considers the KL-divergence as a norm:

$$\nabla_{\boldsymbol{\xi}} \mathcal{L}(\boldsymbol{\xi}) \mathbf{F}_{\boldsymbol{\xi}}^{-1} \propto \lim_{\epsilon \to 0} \frac{1}{\epsilon} \underset{\mathbf{d} \ s.t. \mathsf{KL}[q(\mathbf{u}|\boldsymbol{\xi})|q(\mathbf{u}|\boldsymbol{\xi}+\mathbf{d})] \leq \epsilon}{\mathsf{arg} \ \mathsf{max}} \mathcal{L}(\boldsymbol{\xi} + \mathbf{d})$$

with $\mathbf{F}_{\mathcal{E}}$ the Fisher information of q:

$$\mathbf{F}_{oldsymbol{\xi}} = -\mathbb{E}_{q(\mathbf{u}|oldsymbol{\xi})}[
abla_{oldsymbol{\xi}}^2 \log q(\mathbf{u}|oldsymbol{\xi})]$$

Let η and θ be the natural and expectation parameters of q, respectively:

$$\mathbf{F}_{\eta} = \frac{\partial \boldsymbol{\theta}}{\partial \boldsymbol{\eta}}, \qquad \mathbf{F}_{\xi} = \left(\frac{\partial \boldsymbol{\eta}}{\partial \boldsymbol{\xi}}\right)^{\mathsf{T}} \frac{\partial \boldsymbol{\theta}}{\partial \boldsymbol{\eta}} \frac{\partial \boldsymbol{\eta}}{\partial \boldsymbol{\xi}} = \left(\frac{\partial \boldsymbol{\eta}}{\partial \boldsymbol{\xi}}\right)^{\mathsf{T}} \frac{\partial \boldsymbol{\theta}}{\partial \boldsymbol{\xi}}.$$

Considers the KL-divergence as a norm:

$$abla_{m{\xi}} \mathcal{L}(m{\xi}) \mathbf{F}_{m{\xi}}^{-1} \propto \lim_{\epsilon o 0} rac{1}{\epsilon} rg \max_{\mathbf{d} \, s.t. \mathsf{KL}[q(\mathbf{u}|m{\xi})|q(\mathbf{u}|m{\xi}+\mathbf{d})] \leq \epsilon} \mathcal{L}(m{\xi}+\mathbf{d})$$

with $\mathbf{F}_{\mathcal{E}}$ the Fisher information of q:

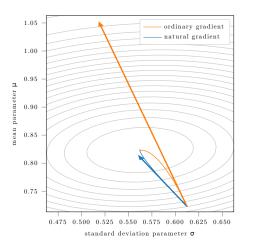
$$\mathbf{F}_{oldsymbol{\xi}} = -\mathbb{E}_{q(\mathbf{u}|oldsymbol{\xi})}[
abla_{oldsymbol{\xi}}^2 \log q(\mathbf{u}|oldsymbol{\xi})]$$

Let η and θ be the natural and expectation parameters of q, respectively:

$$\mathbf{F}_{\boldsymbol{\eta}} = \frac{\partial \boldsymbol{\theta}}{\partial \boldsymbol{\eta}}, \qquad \mathbf{F}_{\boldsymbol{\xi}} = \left(\frac{\partial \boldsymbol{\eta}}{\partial \boldsymbol{\xi}}\right)^{\mathsf{T}} \frac{\partial \boldsymbol{\theta}}{\partial \boldsymbol{\eta}} \frac{\partial \boldsymbol{\eta}}{\partial \boldsymbol{\xi}} = \left(\frac{\partial \boldsymbol{\eta}}{\partial \boldsymbol{\xi}}\right)^{\mathsf{T}} \frac{\partial \boldsymbol{\theta}}{\partial \boldsymbol{\xi}}.$$

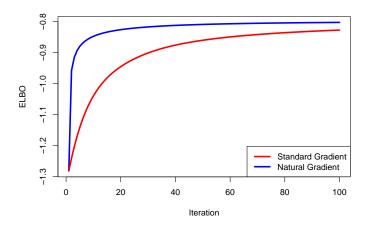
Thus,

$$\nabla_{\boldsymbol{\xi}} \mathcal{L}(\boldsymbol{\xi}) \mathbf{F}_{\boldsymbol{\xi}}^{-1} = \nabla_{\boldsymbol{\xi}} \mathcal{L}(\boldsymbol{\xi}) \frac{\partial \boldsymbol{\xi}}{\partial \boldsymbol{\theta}} \left(\frac{\partial \boldsymbol{\xi}}{\partial \boldsymbol{\eta}} \right)^{\mathsf{T}} = \frac{\partial \mathcal{L}}{\partial \boldsymbol{\theta}} \left(\frac{\partial \boldsymbol{\xi}}{\partial \boldsymbol{\eta}} \right)^{\mathsf{T}}.$$

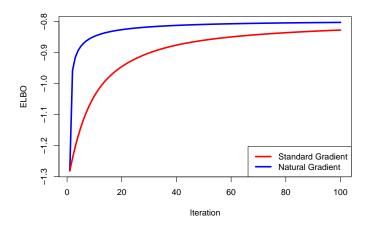


(Salimbeni et al., 2018)

Natural Gradient: Illustrative Example



Natural Gradient: Illustrative Example



The natural gradient achieves a faster convergence!

Run the cells of the notebook to fit a sparse GP using the VFE approximation with natural gradients!

Can we further improve the computational cost in $\mathcal{O}(NM^2)$?

Can we further improve the computational cost in $\mathcal{O}(NM^2)$?

Minibatch training in NN allows to scale to massive datasets!

Can we further improve the computational cost in $\mathcal{O}(NM^2)$?

Minibatch training in NN allows to scale to massive datasets!

Straight forward to do that in the VFE approach:

$$\begin{split} & \mathcal{L}(q, \theta) = \mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f}, \theta)] - \mathsf{KL}[q(\mathbf{u})|p(\mathbf{u})] \\ & = \sum_{i=1}^{N} \mathbb{E}_{q(f_i)}[\log p(y_i|f_i, \theta)] - \mathsf{KL}[q(\mathbf{u})|p(\mathbf{u})] \\ & \approx \frac{N}{B} \sum_{i \in \mathcal{B}} \mathbb{E}_{q(f_i)}[\log p(y_i|f_i, \theta)] - \mathsf{KL}[q(\mathbf{u})|p(\mathbf{u})] \end{split}$$

Can we further improve the computational cost in $\mathcal{O}(NM^2)$?

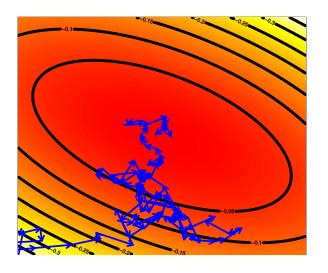
Minibatch training in NN allows to scale to massive datasets!

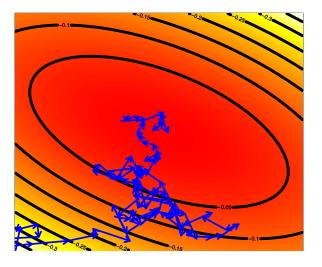
Straight forward to do that in the VFE approach:

$$\begin{split} & \mathcal{L}(q, \theta) = \mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f}, \theta)] - \mathsf{KL}[q(\mathbf{u})|p(\mathbf{u})] \\ & = \sum_{i=1}^{N} \mathbb{E}_{q(f_i)}[\log p(y_i|f_i, \theta)] - \mathsf{KL}[q(\mathbf{u})|p(\mathbf{u})] \\ & \approx \frac{N}{B} \sum_{i \in \mathcal{B}} \mathbb{E}_{q(f_i)}[\log p(y_i|f_i, \theta)] - \mathsf{KL}[q(\mathbf{u})|p(\mathbf{u})] \end{split}$$

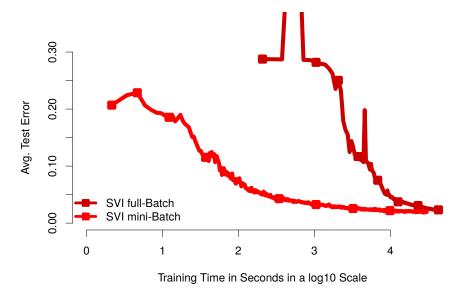
The training cost goes down to $\mathcal{O}(M^3)$ which allows to address datasets with millions of instances!

(Hensman et al., 2013)





To converge to a local neighborhood of the optimum stochastic methods require an estimate of the gradient which can be very cheap!



(Hernández-Lobato, 2015)

• Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- The optimized inducing points spread over the input space where the latent function changes.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- The optimized inducing points spread over the input space where the latent function changes.
- Guaranteed to be exact if M = N and the inducing points are not optimized and located at the training points.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- The optimized inducing points spread over the input space where the latent function changes.
- Guaranteed to be exact if M = N and the inducing points are not optimized and located at the training points.
- It does not change the model. It relies on a particular posterior approximation that speeds-up the computations.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- The optimized inducing points spread over the input space where the latent function changes.
- Guaranteed to be exact if M = N and the inducing points are not optimized and located at the training points.
- It does not change the model. It relies on a particular posterior approximation that speeds-up the computations.
- It allows for minibatch training which reduces the cost to $\mathcal{O}(M^3)$.

- Reduces the cost to $\mathcal{O}(M^2N)$ with $M \ll N$.
- The optimized inducing points spread over the input space where the latent function changes.
- Guaranteed to be exact if M = N and the inducing points are not optimized and located at the training points.
- It does not change the model. It relies on a particular posterior approximation that speeds-up the computations.
- It allows for minibatch training which reduces the cost to $\mathcal{O}(M^3)$.
- The objective is prone to local optima and difficult to optimize.

Run the cells of the notebook to fit a sparse GP using the VFE approximation with mini-batches!

Run the cells of the notebook to fit a sparse GP using the VFE approximation for classification and carry out task 7!

• Exact GPs have an $\mathcal{O}(N^3)$ computational cost, making them feasible on small datasets with a few thousand instances only.

- Exact GPs have an $\mathcal{O}(N^3)$ computational cost, making them feasible on small datasets with a few thousand instances only.
- Sparse GPs provide an approximate solution with a smaller computational cost that is O(NM²) with M ≪ N.

- Exact GPs have an $\mathcal{O}(N^3)$ computational cost, making them feasible on small datasets with a few thousand instances only.
- Sparse GPs provide an approximate solution with a smaller computational cost that is $\mathcal{O}(NM^2)$ with $M \ll N$.
- The non-parametric property of GP is lost when using sparse approximations. They are no longer more flexible with more data.

- Exact GPs have an $\mathcal{O}(N^3)$ computational cost, making them feasible on small datasets with a few thousand instances only.
- Sparse GPs provide an approximate solution with a smaller computational cost that is $\mathcal{O}(NM^2)$ with $M \ll N$.
- The non-parametric property of GP is lost when using sparse approximations. They are no longer more flexible with more data.
- The methods that approximate the GP prior often introduce a low-rank structure in the covariance matrix.

- Exact GPs have an $\mathcal{O}(N^3)$ computational cost, making them feasible on small datasets with a few thousand instances only.
- Sparse GPs provide an approximate solution with a smaller computational cost that is $\mathcal{O}(NM^2)$ with $M \ll N$.
- The non-parametric property of GP is lost when using sparse approximations. They are no longer more flexible with more data.
- The methods that approximate the GP prior often introduce a low-rank structure in the covariance matrix.
- The best performing method seems to be the VFE method since it does not modify the prior.

- Exact GPs have an $\mathcal{O}(N^3)$ computational cost, making them feasible on small datasets with a few thousand instances only.
- Sparse GPs provide an approximate solution with a smaller computational cost that is $\mathcal{O}(NM^2)$ with $M \ll N$.
- The non-parametric property of GP is lost when using sparse approximations. They are no longer more flexible with more data.
- The methods that approximate the GP prior often introduce a low-rank structure in the covariance matrix.
- The best performing method seems to be the VFE method since it does not modify the prior.
- Some methods allow for stochastic optimization and mini-batch training that further reduce the cost to $\mathcal{O}(M^3)$.

References

- Williams, C., & Seeger, M. (2000). Using the Nyström method to speed up kernel machines. Advances in neural information processing systems, 13.
- Snelson, E., & Ghahramani, Z. (2005). Sparse Gaussian processes using pseudo-inputs.
 Advances in neural information processing systems, 18.
- Rahimi, A., & Recht, B. (2007). Random features for large-scale kernel machines.
 Advances in neural information processing systems, 20.
- Naish-Guzman, A., & Holden, S. (2007). The generalized FITC approximation. Advances in neural information processing systems, 20.
- Hernández-Lobato, D., & Hernández-Lobato, J. M. (2016, May). Scalable Gaussian process classification via expectation propagation. In Artificial Intelligence and Statistics (pp. 168-176).
- Hensman, J., Fusi, N., & Lawrence, N. D. (2013). Gaussian processes for big data.
 Proceedings of the Twenty-Ninth Conference on Uncertainty in Artificial Intelligence.
- Salimbeni, H., Eleftheriadis, S., & Hensman, J. (2018, March). Natural gradients in practice: Non-conjugate variational inference in Gaussian process models. In International Conference on Artificial Intelligence and Statistics (pp. 689-697).
- Wu, L., Miller, A., Anderson, L., Pleiss, G., Blei, D., & Cunningham, J. (2021).
 Hierarchical inducing point gaussian process for inter-domian observations.
 International Conference on Artificial Intelligence and Statistics (pp. 2926-2934).