Part II: Sparse Gaussian Processes

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We can handle just a few thousand data instances at most!

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Neural Network (parametric model)

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• VFE: does approximate inference with a simplified distribution q.

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The Woodbury formula gives $(I\sigma^2 + U\Lambda U^T)^{-1}$ with cost $\mathcal{O}(M^2N)!$

$$(A + PCQ)^{-1} = A^{-1} - A^{-1}P (C^{-1} + QA^{-1}P)^{-1} QA^{-1}$$

$$(A + PCQ)^{-1} = A^{-1} - A^{-1}P(C^{-1} + QA^{-1}P)^{-1}QA^{-1}$$

Let us now use $\mathbf{A} = \mathbf{I}\sigma^2$, $\mathbf{P} = \mathbf{U}$, $\mathbf{Q} = \mathbf{U}^{\mathsf{T}}$ and $\mathbf{C} = \mathbf{\Lambda}$.

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Note that A and C are diagonal with sizes $N \times N$ and $M \times M$!

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 $\mathbf{C}^{-1} + \mathbf{Q}\mathbf{A}^{-1}\mathbf{P} = \mathbf{\Lambda}^{-1} + \mathbf{U}^{\mathsf{T}}\mathbf{U}\sigma^{-2}$ has size $M \times M!$

$$(\boldsymbol{\mathsf{A}}+\boldsymbol{\mathsf{PCQ}})^{-1}=\boldsymbol{\mathsf{A}}^{-1}-\boldsymbol{\mathsf{A}}^{-1}\boldsymbol{\mathsf{P}}\left(\boldsymbol{\mathsf{C}}^{-1}+\boldsymbol{\mathsf{Q}}\boldsymbol{\mathsf{A}}^{-1}\boldsymbol{\mathsf{P}}\right)^{-1}\boldsymbol{\mathsf{Q}}\boldsymbol{\mathsf{A}}^{-1}$$

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$$C^{-1} + QA^{-1}P = \Lambda^{-1} + U^{T}U\sigma^{-2}$$
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Computing the whole expression has cost $\mathcal{O}(NM^2)$!

Eigenfunction Analysis of Covariance Functions

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Extended input space: A function $\phi(\cdot)$ that obeys

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is an eigenfunction of $C(\cdot, \cdot)$ with eigenvalue λ , w.r.t., $p(\mathbf{x})$.

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Mercer's theorem:

$$C(\mathbf{x},\mathbf{x}') = \sum_{i=1}^{\infty} \lambda_i \phi_i(\mathbf{x}) \phi_i(\mathbf{x}').$$

An Analytic Example

Consider:

$$p(x) = \mathcal{N}(x|0,\sigma^2), \qquad C(x,x') = \exp\left\{-\frac{1}{2\ell^2}(x-x')^2\right\}.$$

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$$\begin{split} \lambda_k &= \sqrt{\frac{2a}{A}} B^k , \qquad \phi_k(x) = \exp\left\{-(c-a)x^2\right\} H_k(\sqrt{2c}x) ,\\ k &= 0, 1, 2, \dots, \text{ with} \\ &= 4\sigma^2 , \quad b^{-1} = 2\ell^2 , \quad c = \sqrt{a^2 + 2ab} , \quad A = a + b + c , \quad B = b/a , \end{split}$$

and $H_k(\cdot)$, the k-th order Hermite polynomial.

Hermite Polynomials



Covariance Function Approximation

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This motivates the following eigenvalue problem:

$$\lambda_i^{\mathsf{mat}} \mathbf{u}_i = \mathbf{\Sigma} \mathbf{u}_i \,,$$

with $\Sigma_{i,j} = C(\mathbf{x}_i, \mathbf{x}_j)$. Then, we approximate $\phi_i(\mathbf{x}_j) \approx \sqrt{N}(\mathbf{u}_i)_j = \tilde{\phi}_i(\mathbf{x}_j)$, and $\lambda_i \approx \lambda_i^{\text{mat}}/N = \tilde{\lambda}_i$, which guarantees that $\Sigma = \tilde{\Phi} \tilde{\Lambda} \tilde{\Phi}^{\mathsf{T}}$.
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For an arbitrary \mathbf{x}' not in the training set, then:

$$\tilde{\phi}_i(\mathbf{x}') = \frac{1}{N\lambda_i} \sum_{n=1}^N C(\mathbf{x}', \mathbf{x}_n) \phi_i(\mathbf{x}_n) \approx \frac{\sqrt{N}}{\lambda_i^{\text{mat}}} \sum_{n=1}^N C(\mathbf{x}', \mathbf{x}_n) (\mathbf{u}_i)_n = \frac{\sqrt{N}}{\lambda_i^{\text{mat}}} \Sigma(\mathbf{x}')^{\mathsf{T}} \mathbf{u}_i \,.$$

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Using Mercer's theorem and the previous approximation, we approximate the covariance function as:

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which results in a rank M approximation of the covariance matrix Σ :

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The inverse of $I\sigma^2 + \tilde{\Sigma}$ can be efficiently computed using the Woodbury formula with cost $O(NM^2)!$

$$p(\mathbf{f}, \mathbf{f}^{\star}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{f}^{\star} \end{array} \right] \middle| \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array} \right], \left[\begin{array}{c} \tilde{\boldsymbol{\Sigma}} & \boldsymbol{\Sigma}_{\mathbf{f}^{\star}} \\ \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}} & \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} \end{array} \right] \right)$$

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$$\stackrel{\sim}{\xrightarrow{}}_{q} \int_{-\frac{1}{2}}^{\frac{1}{2}} \int_{-\frac{1}{2}}^{\frac{1}{2}$$

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$$\stackrel{\sim}{\xrightarrow{}}_{q} \int_{\mathbf{y}_{1}}^{\mathbf{y}_{1}} \left[\begin{array}{c} \mathbf{0} \\ \mathbf{f}^{\star} \end{array}\right], \left[\begin{array}{c} \mathbf{y}_{1} \\ \mathbf{y}_{2} \end{array}\right], \left[\begin{array}{c} \mathbf{a} \\ \mathbf{b} \end{array}\right], \left[\begin{array}{c} \mathbf{A} & \mathbf{C} \\ \mathbf{C}^{\mathsf{T}} & \mathbf{B} \end{array}\right] \right), \\ p(\mathbf{y}_{1}|\mathbf{y}_{2}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{y}_{1} \\ \mathbf{y}_{2} \end{array}\right], \left[\begin{array}{c} \mathbf{a} \\ \mathbf{b} \end{array}\right], \left[\begin{array}{c} \mathbf{A} & \mathbf{C} \\ \mathbf{C}^{\mathsf{T}} & \mathbf{B} \end{array}\right] \right), \\ p(\mathbf{y}_{1}|\mathbf{y}_{2}) = \mathcal{N}\left(\mathbf{y}_{1}\middle|\mathbf{a} + \mathbf{C}\mathbf{B}^{-1}(\mathbf{y}_{2} - \mathbf{b}), \mathbf{A} - \mathbf{C}\mathbf{B}^{-1}\mathbf{C}^{\mathsf{T}}\right)$$

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$$p(\mathbf{y}_{1}|\mathbf{y}_{2}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{y}_{1} \\ \mathbf{y}_{2} \end{array}\right], \left[\begin{array}{c} \mathbf{a} \\ \mathbf{b} \end{array}\right], \left[\begin{array}{c} \mathbf{A} & \mathbf{C} \\ \mathbf{C}^{\mathsf{T}} & \mathbf{B} \end{array}\right] \right),$$

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$$p(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}| \ \boldsymbol{\Sigma}_{\mathsf{f}^{\star}\mathsf{f}} \mathbf{\tilde{\Sigma}_{\mathsf{f}^{\star}}^{-1}} \mathbf{f}, \boldsymbol{\Sigma}_{\mathsf{f}^{\star}\mathsf{f}^{\star}} - \boldsymbol{\Sigma}_{\mathsf{f}^{\star}\mathsf{f}} \mathbf{\tilde{\Sigma}_{\mathsf{f}^{\star}}^{-1}} \boldsymbol{\Sigma}_{\mathsf{f}^{\star}\mathsf{f}}^{\mathsf{T}}\right)$$

Nyström Approximation: Illustrative Example



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Nystrom GP (M = 10)



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The approximation is similar to the full GP in some regions!

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- It can perform well if Σ is dominated by a few eigenvalues.
- As the *M* points are chosen at random it may give different results.
- Since the approximation is done only over the covariance matrix of the training data, negative predictive variances are possible, but rare.

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Bochner's theorem:

A covariance function $C(\mathbf{x}, \mathbf{x}') = C(\mathbf{x} - \mathbf{x}')$ on \mathbb{R}^D is positive definite if and only if $C(\mathbf{x} - \mathbf{x}')$ is the Fourier transform of a distribution $s(\mathbf{w})$.

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$$C(\mathbf{x}, \mathbf{x}') = \int \exp\{-i\mathbf{w}^{\mathsf{T}}(\mathbf{x} - \mathbf{x}')\}s(\mathbf{w})d\mathbf{w},$$
$$s(\mathbf{w}) = \frac{1}{(2\pi)^{D}}\int \exp\{i\mathbf{w}^{\mathsf{T}}\}C(\tau, \mathbf{0})d\tau.$$

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s(w) is called the spectral density of the covariance function.

Due to Bochner's theorem, the covariance can be written as:

$$C(\mathbf{x}, \mathbf{x}') = \mathbb{E}_{s(\mathbf{w})} \left[\exp\{-i\mathbf{w}^{\mathsf{T}}(\mathbf{x} - \mathbf{x}')\} \right]$$

= $2\mathbb{E}_{s(\mathbf{w}), b \sim U[0, 2\pi]} \left[\cos\left(\mathbf{w}^{\mathsf{T}}\mathbf{x} + b\right) \cos\left(\mathbf{w}^{\mathsf{T}}\mathbf{x}' + b\right) \right]$.

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We can reduce the variance of the estimator by generating M samples:

$$C(\mathbf{x}, \mathbf{x}') \approx \frac{2}{M} \sum_{m=1}^{M} \cos\left(\mathbf{w}_{m}^{\mathsf{T}} \mathbf{x} + b_{m}\right) \cos\left(\mathbf{w}_{m}^{\mathsf{T}} \mathbf{x}' + b_{m}\right) = \phi(\mathbf{x})^{\mathsf{T}} \phi(\mathbf{x}')$$

with $\phi(\mathbf{x}) = \sqrt{\frac{2}{M}} \cos \left(\mathbf{W}^{\mathsf{T}} \mathbf{x} + \mathbf{b} \right)$ a random *M* feature expansion.

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For the squared exponential covariance function s(w) is Gaussian!

Approximate Covariance Function

The covariance matrix can be simply approximated as: $\Sigma\approx\tilde{\Sigma}=\Phi\Phi^{\mathsf{T}}$

and hence $I\sigma^2 + \tilde{\Sigma}$ can be inverted with cost $\mathcal{O}(NM^2)$.

Approximate Covariance Function

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The computational cost is $\mathcal{O}(NM^2)$!

Random Features: Illustrative Example



х
Random Features: Illustrative Example

Random Features GP (M = 10)



Random Features: Illustrative Example

Random Features GP (M = 50)



Random Features: Illustrative Example

Random Features GP (M = 50)



х

In regions with no data the approximation may be wiggling a lot!

• Reduces the cost to $\mathcal{O}(MN^2)$ with $M \ll N$.

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- Very simple to implement!
- Equivalent to a neural network with a hidden layer with *M* units and cosine activations, and a Bayesian linear model in the last layer!

1. Extend model with $M \ll N$ inducing points and outputs at $\overline{\mathbf{X}}$.

$$\rho(\mathbf{f},\mathbf{u}) = \mathcal{N}\left(\left[\begin{array}{c} \mathbf{f} \\ \mathbf{u} \end{array} \right] \middle| \left[\begin{array}{c} \mathbf{0} \\ \mathbf{0} \end{array} \right], \left[\begin{array}{c} \boldsymbol{\Sigma}_{\mathrm{ff}} & \boldsymbol{\Sigma}_{\mathrm{fu}} \\ \boldsymbol{\Sigma}_{\mathrm{uf}} & \boldsymbol{\Sigma}_{\mathrm{uu}} \end{array} \right] \right)$$

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2. Introduce conditional independences:

$$p(\mathbf{f}|\mathbf{u}) = \prod_{i=1}^{N} p(f_i|\mathbf{u})$$

3. Marginalize \mathbf{u} to obtain an approximate GP prior for \mathbf{f} .

$$p(\mathbf{f}) = \int p(\mathbf{f}|\mathbf{u})p(\mathbf{u})d\mathbf{u} = \prod_{i=1}^{N} p(f_i|\mathbf{u})p(\mathbf{u})d\mathbf{u} = \mathcal{N}(\mathbf{f}|0, \tilde{\boldsymbol{\Sigma}}_{\mathbf{ff}})$$

where $\tilde{\Sigma}_{ff} = \mathbf{D} + \mathbf{Q}_{ff}$ with **D** diagonal and $\mathbf{Q}_{ff} = \Sigma_{fu} \Sigma_{uu}^{-1} \Sigma_{uf}$ of rank *M*.

5. We make the prediction of f^* at \mathbf{x}^* by considering the approximate GP prior:

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$$(1)$$
 (1)



$$\begin{split} p(\mathbf{y}_1, \mathbf{y}_2) &= \mathcal{N}\left(\begin{bmatrix} \mathbf{y}_1 \\ \mathbf{y}_2 \end{bmatrix}, \begin{bmatrix} \mathbf{a} \\ \mathbf{b} \end{bmatrix}, \begin{bmatrix} \mathbf{A} & \mathbf{C} \\ \mathbf{C}^{\mathsf{T}} & \mathbf{B} \end{bmatrix} \right) \,, \\ p(\mathbf{y}_1 | \mathbf{y}_2) &= \frac{p(\mathbf{y}_1, \mathbf{y}_2)}{p(\mathbf{y}_2)} \,, \end{split}$$

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 u_1

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6. How do we find the location of the inducing points $\overline{\mathbf{X}}$?

Simply treat them as prior parameters and maximize the approximate marginal likelihood $p(f|0, \tilde{\Sigma}_{ff})!$

FITC: Illustrative Example

Full GP



х

FITC: Illustrative Example

FITC (M = 10)



х

FITC: Illustrative Example

FITC (M = 10)



The inducing points cover the regions where the function changes!

• Reduces the cost to $\mathcal{O}(MN^2)$ with $M \ll N$.

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- It can be understood as considering heteroscedastic (input dependent) noise!

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 $p(y_i|f(\mathbf{x}_i)) = \phi(y_i f(\mathbf{x}_i)), \quad \phi(\cdot) \equiv \text{The c.d.f. of a standard Gaussian.}$

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Approximates with a Gaussian distribution the intractable posterior:

$$p(\mathbf{f}|\mathbf{y}) = rac{\prod_{i=1}^N \phi(y_i f(\mathbf{x}_i)) \mathcal{N}(\mathbf{f}|\mathbf{0}, ilde{\mathbf{\Sigma}})}{p(\mathbf{y})}\,,$$

where $\tilde{\Sigma}$ is the approximate FITC covariance matrix.

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The exponential family:

$$q(\mathbf{x}) = \exp\left(\eta^{\mathsf{T}}\mathbf{u}(\mathbf{x}) - g(\eta)
ight), \quad g(\eta) = \log\int\exp\left(\eta^{\mathsf{T}}\mathbf{u}(\mathbf{x})
ight)d\mathbf{x}$$

where η is a vector of natural parameters of q, $\mathbf{u}(\mathbf{x})$ are the sufficient statistics and $g(\eta)$ is a **log partition function**.

Examples of Distributions in the Exponential Family

Gaussian:

$$\mathcal{N}(x|\mu,\sigma^2) = 1/\sqrt{2\pi\sigma^2} \exp\left\{-\frac{1}{2\sigma^2}(x-\mu)^2\right\}$$
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Most parametric distributions belong to the exponential family!

$$p_1(x) = \frac{1}{\sqrt{2\pi\sigma_1^2}} \exp\left\{-\frac{1}{2}\sigma_1^2(x-\mu_1)^2\right\} ,$$

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Consider these two Gaussian distributions:

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Consider the KL-divergence between p and q (q in the exponential family):

$$\mathsf{KL}(p||q) = -\int p(\mathbf{x}) \log \left\{ \frac{q(\mathbf{x})}{p(\mathbf{x})} \right\} d\mathbf{x} = g(\eta) - \eta^{\mathsf{T}} \mathbb{E}_p[\mathbf{u}(\mathbf{x})] + \mathsf{Const} \,.$$

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KL(p||q) is minimized by matching expected sufficient statistics.

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Furthermore, it is possible to show that:

$$rac{\partial g(\boldsymbol{\eta})}{\partial \boldsymbol{\eta}} = \mathbb{E}_{\boldsymbol{q}}[\boldsymbol{\mathsf{u}}(\boldsymbol{\mathsf{x}})]\,.$$

 $\mathsf{KL}(p||q)$ is minimized by matching expected sufficient statistics. If q is Gaussian, then we have to match $\mathbb{E}_q[\mathbf{x}] = \mathbb{E}_p[\mathbf{x}]$ and $\mathbb{E}_q[\mathbf{x}\mathbf{x}^\mathsf{T}] = \mathbb{E}_p[\mathbf{x}\mathbf{x}^\mathsf{T}].$

Joint Approximation

EP approximates this joint distribution by a product of simpler factors:

$$p(\mathbf{f},\mathbf{y}) = \prod_{i=1}^{N} \phi_i(y_i f(\mathbf{x}_i) \mathcal{N}(\mathbf{f}|\mathbf{0}, \tilde{\boldsymbol{\Sigma}}) = \prod_i t_i(\mathbf{f}) \approx \prod_i \tilde{t}_i(\mathbf{f}),$$

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The exponential family is closed under the product and $\prod_i \tilde{t}_i$ can be easily normalized to compute an approximate distribution:

$$p(\mathbf{f}|\mathbf{y}) = \frac{1}{p(\mathbf{y})} \prod_i t_i(\mathbf{f}) \approx \frac{1}{Z} \prod_i \tilde{t}_i(\mathbf{f}) = q(\mathbf{f}),$$

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Therefore q has the same form as the approximate factors!

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Suppose we wish to refine \tilde{t}_j . We first remove this factor from the product:

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Then, \tilde{t}_j is updated to minimize the KL-divergence between:

where q^{i} is fixed. This ensures that \tilde{t}_j is accurate where q^{i} is high.

In practice, \tilde{t}_i is found by first **minimizing** with respect to q_{new} :

$$\mathsf{KL}\left(\left.\frac{t_j(\mathbf{f})q^{\setminus j}(\mathbf{f})}{Z_j}\right|q_{\mathsf{new}}(\mathbf{f})
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The refined factor \tilde{t}_j is set in practice to be:

$$ilde{t}_j(\mathbf{f}) = Z_j rac{q_{\mathsf{new}}(\mathbf{f})}{q^{\setminus j}(\mathbf{f})}\,, \qquad \qquad ext{with} \quad ilde{t}_j(\mathbf{f}) q^{\setminus j}(\mathbf{f}) \propto q_{\mathsf{new}}\,,$$

which ensures that $\tilde{t}_j(\mathbf{f})q^{ij}(\mathbf{f})$ and $t_j(\mathbf{f})q^{ij}(\mathbf{f})$ integrate the same.

Several passes are made trough the factors until they **converge**. The **model evidence** is approximated by the **normalizing constant** of *q*.

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The FITC prior results in a total cost of $\mathcal{O}(NM^2)$!

Graphical Illustration

Approximates
$$p(\mathbf{f}|\mathbf{y}) \propto t_0(\mathbf{f}) \prod_{j=1}^N t_j(\mathbf{f})$$
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$$p(\mathbf{f}|\mathbf{y}) \propto t_0(\mathbf{f}) \quad t_1(\mathbf{f}) \quad t_2(\mathbf{f}) \quad t_3(\mathbf{f}) \approx q(\mathbf{f}) \propto t_0(\mathbf{f}) \quad \tilde{t}_1(\mathbf{f}) \quad \tilde{t}_2(\mathbf{f}) \quad \tilde{t}_3(\mathbf{f})$$

The \tilde{t}_j are tuned by minimizing the KL-divergence

$$\mathsf{KL}[\hat{p}_j||q] \quad \text{for}\, j = 1, \dots, N\,, \quad \text{where} \quad \begin{array}{l} \hat{p}_j(\mathbf{f}) \quad \propto \quad t_j(\mathbf{f}) \prod_{i \neq j} \tilde{t}_i(\mathbf{f}) \\ q(\mathbf{f}) \quad \propto \quad \tilde{t}_j(\mathbf{f}) \prod_{i \neq j} \tilde{t}_i(\mathbf{f})\,. \end{array}$$

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If the exact factor already belongs to the exponential family it needs not be approximated!

GFITC: Factor Approximation



f(x_i)

GFITC: Factor Approximation



The approximate factor is accurate in regions of high posterior probability as indicated by the cavity distribution!

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- Direct dependence of log Z on the hyper-parameters.
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If EP converges the gradient of $\log Z$ w.r.t. each θ_i is zero, which allows to easily compute the gradients of $\log Z$!

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$$p(\mathbf{f}^{\star}, \mathbf{f}) = \mathcal{N}\left(\begin{bmatrix} \mathbf{f}^{\star} \\ \mathbf{f} \end{bmatrix} \middle| \begin{bmatrix} \mathbf{0} \\ \mathbf{0} \end{bmatrix}, \begin{bmatrix} \mathbf{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} & \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}} \\ \mathbf{Q}_{\mathbf{f}\mathbf{f}^{\star}} & \tilde{\boldsymbol{\Sigma}_{\mathbf{f}^{\star}}} \end{bmatrix} \right)$$

We want to compute the value of f^* at a new x^* :

$$p(\mathbf{f}^{\star},\mathbf{f}) = \mathcal{N}\left(\left[\begin{array}{c}\mathbf{f}^{\star}\\\mathbf{f}\end{array}\right] \middle| \left[\begin{array}{c}\mathbf{0}\\\mathbf{0}\end{array}\right], \left[\begin{array}{c}\boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} & \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\\\mathbf{Q}_{\mathbf{f}\mathbf{f}^{\star}} & \boldsymbol{\tilde{\Sigma}_{\mathbf{f}^{\star}}}\end{array}\right]\right)$$

The conditional $p(\mathbf{f}^*|\mathbf{f})$ is:

$$p(\mathbf{f}^{\star}|\mathbf{f}) = \mathcal{N}\left(\mathbf{f}^{\star}| \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}} \tilde{\boldsymbol{\Sigma}}_{\mathbf{f}}^{-1} \mathbf{f}, \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}} \tilde{\boldsymbol{\Sigma}_{\mathbf{f}}}^{-1} \mathbf{Q}_{\mathbf{f}^{\star}\mathbf{f}}\right)$$

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After marginalizing \mathbf{f} w.r.t. $q(\mathbf{f})$, we obtain the predictive distribution:

$$\begin{split} p(\mathbf{f}^{\star}|\mathbf{y}) &= \int p(\mathbf{f}^{\star}|\mathbf{f})q(\mathbf{f})d\mathbf{f} \\ &= \mathcal{N}\left(\mathbf{f}^{\star}|\;\mathbf{Q}_{\mathsf{f}^{\star}\mathsf{f}}\tilde{\boldsymbol{\Sigma}}_{\mathsf{ff}}^{-1}\tilde{\mathbf{y}},\boldsymbol{\Sigma}_{\mathsf{f}^{\star}\mathsf{f}^{\star}} - \mathbf{Q}_{\mathsf{f}^{\star}\mathsf{f}}\left(\tilde{\boldsymbol{\Sigma}}_{\mathsf{ff}} + \tilde{\boldsymbol{\Pi}}\right)^{-1}\mathbf{Q}_{\mathsf{ff}^{\star}}\right) \end{split}$$

GFITC: Illustrative Example

Full GP + EP



x1

GFITC: Illustrative Example

GFITC (M=10)



x1

GFITC: Illustrative Example

GFITC (M=10)



The inducing points spread across the input space!

Previous methods approximate the GP prior using a low rank approximation of Σ , resulting in a cost $\mathcal{O}(NM^2)$.

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Since the GP prior is not changed it tends to perform better than the previous methods!

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The expression for the KL divergence between q and p is:

$$\int q(\mathbf{f}) \log \frac{q(\mathbf{f})}{p(\mathbf{f})} d\mathbf{f} \geq 0$$

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KL(q|p) depends on p, which is assumed to be intractable! Let the target be p(f|y). Consider the decomposition of p(y):

$$\log p(\mathbf{y}) = \mathcal{L}(q) + \mathsf{KL}(q|p),$$

where

$$\mathcal{L}(q) = \int q(\mathbf{f}) \log rac{p(\mathbf{f}, \mathbf{y})}{q(\mathbf{f})} d\mathbf{f}, \quad \mathsf{KL}(q|p) = \int q(\mathbf{f}) \log rac{q(\mathbf{f})}{p(\mathbf{f}|\mathbf{y})} d\mathbf{f}.$$

Decomposition of the Marginal Likelihood



Decomposition of the Marginal Likelihood



 $\mathcal{L}(q)$ can be used to approximate $\log p(\mathbf{y})$ if $\mathsf{KL}(q|p)$ is small!

Lower bound the log-likelihood:

$$\log p(\mathbf{y}|\theta) = \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u}|\theta) d\mathbf{f} d\mathbf{u}$$



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$$= \log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u} | \theta) \frac{q(\mathbf{f}, \mathbf{u})}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u} \geq \int q(\mathbf{f}, \mathbf{u}) \log \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u} | \theta)}{q(\mathbf{f}, \mathbf{u})} d\mathbf{f} d\mathbf{u} \equiv \mathcal{L}(q, \theta)$$

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$$\mathcal{L}(q,\theta) = \int q(\mathbf{f},\mathbf{u}) \log \frac{p(\mathbf{y},\mathbf{f},\mathbf{u}|\theta)}{q(\mathbf{f},\mathbf{u})} d\mathbf{f} d\mathbf{u} = \log p(\mathbf{y}|\theta) - \mathsf{KL}[q(\mathbf{f},\mathbf{u})|p(\mathbf{f},\mathbf{u}|\mathbf{y})]$$



 $KL \equiv Kullback-Leibler divergence$



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By maximizing $\mathcal{L}(q, \theta)$ w.r.t q we are enforcing that $q(\mathbf{f}, \mathbf{u})$ looks similar to $p(\mathbf{f}, \mathbf{u}|\mathbf{y})$ in terms of the KL!
Consider the following approximate distribution:

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Consider the following approximate distribution:



Consider the following approximate distribution:



The inducing points are now parameters of the approx. dist. q!

$$\begin{aligned} \mathcal{L}(q,\theta) &= \int q(\mathbf{f},\mathbf{u}) \log \frac{p(\mathbf{y},\mathbf{f},\mathbf{u}|\theta)}{q(\mathbf{f},\mathbf{u})} d\mathbf{f} d\mathbf{u} \\ &= \int p(\mathbf{f}|\mathbf{u})q(\mathbf{u}) \log \frac{p(\mathbf{y}|\mathbf{f},\theta)p(\mathbf{f}|\mathbf{u})p(\mathbf{u})}{p(\mathbf{f}|\mathbf{u})q(\mathbf{u})} d\mathbf{f} d\mathbf{u} \end{aligned}$$

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• K

$$\begin{aligned} \mathcal{L}(q,\theta) &= \int q(\mathbf{f},\mathbf{u}) \log \frac{p(\mathbf{y},\mathbf{f},\mathbf{u}|\theta)}{q(\mathbf{f},\mathbf{u})} d\mathbf{f} d\mathbf{u} \\ &= \int p(\mathbf{f}|\mathbf{u})q(\mathbf{u}) \log \frac{p(\mathbf{y}|\mathbf{f},\theta)p(\mathbf{f}|\mathbf{u})p(\mathbf{u})}{p(\mathbf{f}|\mathbf{u})q(\mathbf{u})} d\mathbf{f} d\mathbf{u} \end{aligned}$$

$$\mathcal{L}(q, \theta) = \mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f}, \theta)] - \mathsf{KL}[q(\mathbf{u})|p(\mathbf{u})]$$
• Mean squared prediction error \hat{f}
• KL between Gaussians

Plugging $q(\mathbf{f}, \mathbf{u})$ into the lower bound we have:

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Mea

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- No change in the model is made and the cost is in $\mathcal{O}(M^2N)$!
- Predictions are made using $p(\mathbf{f}^*|\mathbf{u})q(\mathbf{u})$ marginalizing out \mathbf{u} .

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We want to compute the value of f^* at a new x^* :

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The conditional $p(\mathbf{f}^*|\mathbf{u})$ is:

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After marginalizing **u** w.r.t. $q(\mathbf{u})$, we obtain the predictive distribution:

$$p(\mathbf{f}^{\star}|\mathbf{y}) = \int p(\mathbf{f}^{\star}|\mathbf{u})q(\mathbf{u})d\mathbf{u}$$
$$= \mathcal{N}(\mathbf{f}^{\star}|\boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{u}}\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}^{-1}\mathbf{m}, \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{f}^{\star}} - \boldsymbol{\Sigma}_{\mathbf{f}^{\star}\mathbf{u}} \left(\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}^{-1} - \boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}^{-1}\mathbf{S}\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{u}}^{-1}\right)\boldsymbol{\Sigma}_{\mathbf{u}\mathbf{f}^{\star}})$$

VFE: Illustrative Example

Full GP



х

VFE: Illustrative Example

VFE (M = 10)



х

VFE: Illustrative Example

VFE (M = 10)



The inducing points cover the regions where the function changes!

VFE: Illustrative Classification Example

Full GP + EP



x1

VFE: Illustrative Classification Example

VFE (M=10)



x1

VFE: Illustrative Classification Example

VFE (M=10)



The inducing points spread across the input space!

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- FITC: less local optima and easier to optimize, also less accurate.
- VFE: more accurate, more local optima, more difficult to optimize.

(Bui et al., 2017) (Bauer et al., 2016)

Whitened Parameterization for VFE

Alternative VFE objective expected to be easier to optimize!

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with **u** the latent process values at the inducing points and $\mathbf{L}^{\mathsf{T}}\mathbf{L} = \Sigma_{uu}$.

The VFE objective becomes:

$$\sum_{i=1}^{N} \mathbb{E}_{q(\mathbf{e})p(f(\mathbf{x}_i)|\mathbf{e})} \left[\log p(y_i|f(\mathbf{x}_i)) \right] - \mathsf{KL}(q(\mathbf{e})|\mathcal{N}(\mathbf{0},\mathbf{I})) ,$$

with $p(f(\mathbf{x}_i)|\mathbf{e})$ using the covariances between $f(\mathbf{x}_i)$ and \mathbf{e} .

Whitened Parameterization: Illustrative Example



Whitened Parameterization: Illustrative Example



Whitening significantly improves convergence!

Gradient ascent moves in the direction of the gradient $\nabla_{\boldsymbol{\xi}} \mathcal{L}(\boldsymbol{\xi})$.

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$$abla_{oldsymbol{\xi}} \mathcal{L}(oldsymbol{\xi}) \propto \lim_{\epsilon o 0} rac{1}{\epsilon} rgmax_{oldsymbol{d}, s.t. ||oldsymbol{d}|| \leq \epsilon} \mathcal{L}(oldsymbol{\xi} + \epsilon oldsymbol{d})$$

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If $\boldsymbol{\xi}$ represents the parameters of probability distributions, the Euclidean norm may be problematic!

Illustration with Two Gaussians



х

Illustration with Two Gaussians



х

Illustration with Two Gaussians



The Euclidean distance between parameters is 4 in both cases!
Illustration with Two Gaussians



A better alternative is the KL-divergence between distributions!

Considers the KL-divergence as a norm:

$$\nabla_{\boldsymbol{\xi}} \mathcal{L}(\boldsymbol{\xi}) \mathbf{F}_{\boldsymbol{\xi}}^{-1} \propto \lim_{\epsilon \to 0} \frac{1}{\epsilon} \max_{\mathbf{d} \ \boldsymbol{s}. t. \mathsf{KL}[q(\mathbf{u}|\boldsymbol{\xi})]q(\mathbf{u}|\boldsymbol{\xi}+\mathbf{d})] \leq \epsilon} \mathcal{L}(\boldsymbol{\xi} + \epsilon \mathbf{d})$$

with $\mathbf{F}_{\boldsymbol{\xi}}$ the Fisher information of q:

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Let η and θ be the natural and expectation parameters of q, respectively:

$$\mathbf{F}_{\boldsymbol{\eta}} = \frac{\partial \boldsymbol{\theta}}{\partial \boldsymbol{\eta}}, \qquad \qquad \mathbf{F}_{\boldsymbol{\xi}} = \left(\frac{\partial \boldsymbol{\eta}}{\partial \boldsymbol{\xi}}\right)^{\mathsf{T}} \frac{\partial \boldsymbol{\theta}}{\partial \boldsymbol{\eta}} \frac{\partial \boldsymbol{\eta}}{\partial \boldsymbol{\xi}}.$$

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Thus,

$$\nabla_{\boldsymbol{\xi}} \mathcal{L}(\boldsymbol{\xi}) \mathbf{F}_{\boldsymbol{\xi}}^{-1} = \frac{\partial \mathcal{L}}{\partial \boldsymbol{\theta}} \left(\frac{\partial \boldsymbol{\xi}}{\partial \boldsymbol{\eta}} \right)^{\mathsf{T}}$$

•



(Salimbeni et al., 2018)

Natural Gradient: Illustrative Example



Natural Gradient: Illustrative Example



The natural gradient achieves a faster convergence!

Can we further improve the computational cost in $\mathcal{O}(NM^2)$?

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Straight forward to do that in the VFE approach:

$$\begin{split} \mathcal{L}(\boldsymbol{q}, \boldsymbol{\theta}) &= \mathbb{E}_{\boldsymbol{q}(\mathbf{f})}[\log p(\mathbf{y}|\mathbf{f}, \boldsymbol{\theta})] - \mathsf{KL}[\boldsymbol{q}(\mathbf{u})|\boldsymbol{p}(\mathbf{u})] \\ &= \sum_{i=1}^{N} \mathbb{E}_{\boldsymbol{q}(f_i)}[\log p(y_i|f_i, \boldsymbol{\theta})] - \mathsf{KL}[\boldsymbol{q}(\mathbf{u})|\boldsymbol{p}(\mathbf{u})] \\ &\approx \frac{B}{N} \sum_{i \in \mathcal{B}} \mathbb{E}_{\boldsymbol{q}(f_i)}[\log p(y_i|f_i, \boldsymbol{\theta})] - \mathsf{KL}[\boldsymbol{q}(\mathbf{u})|\boldsymbol{p}(\mathbf{u})] \end{split}$$

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The training cost goes down to $\mathcal{O}(M^3)$ which allows to address datasets with millions of instances!

(Hensman et al., 2013)





To converge to a local neighborhood of the optimum stochastic methods require an estimate of the gradient which can be very cheap!



Training Time in Seconds in a log10 Scale

(Hernández-Lobato, 2015)

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- It does not change the model. It relies on a particular posterior approximation that speeds-up the computations.
- It allows for minibatch training which reduces the cost to $\mathcal{O}(M^3)$.
- The objective is prone to local optima and difficult to optimize.

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- The methods that approximate the GP prior often introduce a low-rank structure in the covariance matrix.
- The best performing method seems to be the VFE method since it does not modify the prior.
- Some methods allow for stochastic optimization and mini-batch training that further reduce the cost to $\mathcal{O}(M^3)$.

References

- Williams, C., & Seeger, M. (2000). Using the Nyström method to speed up kernel machines. Advances in neural information processing systems, 13.
- Snelson, E., & Ghahramani, Z. (2005). Sparse Gaussian processes using pseudo-inputs. Advances in neural information processing systems, 18.
- Rahimi, A., & Recht, B. (2007). Random features for large-scale kernel machines. Advances in neural information processing systems, 20.
- Naish-Guzman, A., & Holden, S. (2007). The generalized FITC approximation. Advances in neural information processing systems, 20.
- Hernández-Lobato, D., & Hernández-Lobato, J. M. (2016, May). Scalable Gaussian process classification via expectation propagation. In Artificial Intelligence and Statistics (pp. 168-176).
- Hensman, J., Fusi, N., & Lawrence, N. D. (2013). Gaussian processes for big data. Proceedings of the Twenty-Ninth Conference on Uncertainty in Artificial Intelligence.
- Salimbeni, H., Eleftheriadis, S., & Hensman, J. (2018, March). Natural gradients in practice: Non-conjugate variational inference in Gaussian process models. In International Conference on Artificial Intelligence and Statistics (pp. 689-697).
- Wu, L., Miller, A., Anderson, L., Pleiss, G., Blei, D., & Cunningham, J. (2021). Hierarchical inducing point gaussian process for inter-domian observations. International Conference on Artificial Intelligence and Statistics (pp. 2926-2934).