# Part II: Sparse Gaussian Processes 

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We can handle just a few thousand data instances at most!

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Neural Network (parametric model)

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- Nyström, Random Features and FITC: approximate GP prior!
- VFE: does approximate inference with a simplified distribution $q$.


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A low rank $m$ approximation of $\boldsymbol{\Sigma}$ does the job:


The Woodbury formula gives $\left(\mathbf{I} \sigma^{2}+\mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^{\mathbf{T}}\right)^{-1}$ with cost $\mathcal{O}\left(M^{2} N\right)$ !

## Woodbury Formula

$$
(\mathbf{A}+\mathbf{P C Q})^{-1}=\mathbf{A}^{-1}-\mathbf{A}^{-1} \mathbf{P}\left(\mathbf{C}^{-1}+\mathbf{Q A}^{-1} \mathbf{P}\right)^{-1} \mathbf{\mathbf { Q A } ^ { - 1 }}
$$

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Let us now use $\mathbf{A}=\mathbf{I} \sigma^{2}, \mathbf{P}=\mathbf{U}, \mathbf{Q}=\mathbf{U}^{\top}$ and $\mathbf{C}=\boldsymbol{\Lambda}$.

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\mathbf{C}^{-1}+\mathbf{Q A}^{-1} \mathbf{P}=\mathbf{\Lambda}^{-1}+\mathbf{U}^{\top} \mathbf{U} \sigma^{-2} \text { has size } M \times M!
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$\mathbf{C}^{-1}+\mathbf{Q A}^{-1} \mathbf{P}=\boldsymbol{\Lambda}^{-1}+\mathbf{U}^{\top} \mathbf{U} \sigma^{-2}$ has size $M \times M!$

$$
\left(\mathbf{I} \sigma^{2}+\mathbf{U} \boldsymbol{\Lambda} \mathbf{U}^{\top}\right)^{-1}=\mathbf{I} \sigma^{-2}-\sigma^{-2} \mathbf{U}\left(\boldsymbol{\Lambda}^{-1}+\mathbf{U}^{\top} \sigma^{-2} \mathbf{U}\right)^{-1} \mathbf{U}^{\top} \sigma^{-2}
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$$

Computing the whole expression has cost $\mathcal{O}\left(N M^{2}\right)$ !

## Eigenfunction Analysis of Covariance Functions

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Extended input space: A function $\phi(\cdot)$ that obeys

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Mercer's theorem:

$$
C\left(\mathbf{x}, \mathbf{x}^{\prime}\right)=\sum_{i=1}^{\infty} \lambda_{i} \phi_{i}(\mathbf{x}) \phi_{i}\left(\mathbf{x}^{\prime}\right)
$$

## An Analytic Example

Consider:

$$
p(x)=\mathcal{N}\left(x \mid 0, \sigma^{2}\right), \quad C\left(x, x^{\prime}\right)=\exp \left\{-\frac{1}{2 \ell^{2}}\left(x-x^{\prime}\right)^{2}\right\}
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$$

Then,

$$
\lambda_{k}=\sqrt{\frac{2 a}{A}} B^{k}, \quad \phi_{k}(x)=\exp \left\{-(c-a) x^{2}\right\} H_{k}(\sqrt{2 c} x)
$$

for $k=0,1,2, \ldots$, with
$a^{-1}=4 \sigma^{2}, \quad b^{-1}=2 \ell^{2}, \quad c=\sqrt{a^{2}+2 a b}, \quad A=a+b+c, \quad B=b / a$,
and $H_{k}(\cdot)$, the k-th order Hermite polynomial.

## Hermite Polynomials



## Covariance Function Approximation

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Exact Covariance Matrix


Approx. Covariance Matrix


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This motivates the following eigenvalue problem:

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\lambda_{i}^{\mathrm{mat}} \mathbf{u}_{i}=\boldsymbol{\Sigma} \mathbf{u}_{i}
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with $\Sigma_{i, j}=C\left(\mathbf{x}_{i}, \mathbf{x}_{j}\right)_{\tilde{\lambda}_{i}}$ Then, we approximate $\phi_{i}\left(\mathbf{x}_{j}\right) \approx \sqrt{N}\left(\mathbf{u}_{i}\right)_{j}=\tilde{\phi}_{i}\left(\mathbf{x}_{j}\right)$, and $\lambda_{i} \approx \lambda_{i}^{\text {mat }} / N=\tilde{\lambda}_{i}$, which guarantees that $\Sigma=\tilde{\Phi} \tilde{\Lambda} \tilde{\Phi}^{\top}$.

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For an arbitrary $\mathbf{x}^{\prime}$ not in the training set, then:

$$
\tilde{\phi}_{i}\left(\mathbf{x}^{\prime}\right)=\frac{1}{N \lambda_{i}} \sum_{n=1}^{N} C\left(\mathbf{x}^{\prime}, \mathbf{x}_{n}\right) \phi_{i}\left(\mathbf{x}_{n}\right) \approx \frac{\sqrt{N}}{\lambda_{i}^{\text {mat }}} \sum_{n=1}^{N} C\left(\mathbf{x}^{\prime}, \mathbf{x}_{n}\right)\left(\mathbf{u}_{i}\right)_{n}=\frac{\sqrt{N}}{\lambda_{i}^{\text {mat }}} \boldsymbol{\Sigma}\left(\mathbf{x}^{\prime}\right)^{\top} \mathbf{u}_{i}
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## Putting All Together

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Using Mercer's theorem and the previous approximation, we approximate the covariance function as:

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which results in a rank $M$ approximation of the covariance matrix $\boldsymbol{\Sigma}$ :

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\boldsymbol{\Sigma} \approx \tilde{\boldsymbol{\Sigma}}=\tilde{\boldsymbol{\Phi}} \tilde{\boldsymbol{\Lambda}} \tilde{\boldsymbol{\Phi}}^{\top}=\boldsymbol{\Sigma}_{N, M} \boldsymbol{\Sigma}_{M, M}^{-1} \boldsymbol{\Sigma}_{M, N}
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The inverse of $\mathbf{I} \sigma^{2}+\tilde{\boldsymbol{\Sigma}}$ can be efficiently computed using the Woodbury formula with cost $\mathcal{O}\left(N M^{2}\right)$ !

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\mathbf{0}
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\end{array}\right]\right), \\
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p\left(\mathbf{f}^{\star} \mid \mathbf{f}\right)=\mathcal{N}\left(\mathbf{f}^{\star} \mid \boldsymbol{\Sigma}_{\mathbf{f}_{\star}} \tilde{\boldsymbol{\Sigma}}_{\mathbf{f f}}^{-1} \mathbf{f}, \boldsymbol{\Sigma}_{\mathbf{f}^{\star} \mathbf{f}^{\star}}-\boldsymbol{\Sigma}_{\mathbf{f} \star \mathbf{f}} \tilde{\Sigma}_{\mathbf{f f}}^{-1} \boldsymbol{\Sigma}_{\mathbf{f} \neq \mathbf{f}}^{\top}\right)
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## Nyström Approximation: Illustrative Example

Full GP


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Nystrom GP (M = 10)


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The approximation is similar to the full GP in some regions!

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- It can perform well if $\boldsymbol{\Sigma}$ is dominated by a few eigenvalues.
- As the $M$ points are chosen at random it may give different results.
- Since the approximation is done only over the covariance matrix of the training data, negative predictive variances are possible, but rare.


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A covariance function $C\left(\mathbf{x}, \mathbf{x}^{\prime}\right)=C\left(\mathbf{x}-\mathbf{x}^{\prime}\right)$ on $\mathbb{R}^{D}$ is positive definite if and only if $C\left(\mathbf{x}-\mathbf{x}^{\prime}\right)$ is the Fourier transform of a distribution $s(\mathbf{w})$.

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$$
\begin{aligned}
C\left(\mathbf{x}, \mathbf{x}^{\prime}\right) & =\int \exp \left\{-i \mathbf{w}^{\top}\left(\mathbf{x}-\mathbf{x}^{\prime}\right)\right\} s(\mathbf{w}) d \mathbf{w} \\
s(\mathbf{w}) & =\frac{1}{(2 \pi)^{D}} \int \exp \left\{i \mathbf{w}^{\top}\right\} C(\boldsymbol{\tau}, \mathbf{0}) d \boldsymbol{\tau}
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They can be used to approximate any stationary covariance function (only depends on the distance between points).

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A covariance function $C\left(\mathbf{x}, \mathbf{x}^{\prime}\right)=C\left(\mathbf{x}-\mathbf{x}^{\prime}\right)$ on $\mathbb{R}^{D}$ is positive definite if and only if $C\left(\mathbf{x}-\mathbf{x}^{\prime}\right)$ is the Fourier transform of a distribution $s(\mathbf{w})$.

$$
\begin{aligned}
C\left(\mathbf{x}, \mathbf{x}^{\prime}\right) & =\int \exp \left\{-i \mathbf{w}^{\top}\left(\mathbf{x}-\mathbf{x}^{\prime}\right)\right\} s(\mathbf{w}) d \mathbf{w} \\
s(\mathbf{w}) & =\frac{1}{(2 \pi)^{D}} \int \exp \left\{i \mathbf{w}^{\top}\right\} C(\boldsymbol{\tau}, \mathbf{0}) d \boldsymbol{\tau} .
\end{aligned}
$$

$s(\mathbf{w})$ is called the spectral density of the covariance function.

## Covariances as Expectations of Cosines

Due to Bochner's theorem, the covariance can be written as:

$$
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C\left(\mathbf{x}, \mathbf{x}^{\prime}\right) & =\mathbb{E}_{s(\mathbf{w})}\left[\exp \left\{-i \mathbf{w}^{\top}\left(\mathbf{x}-\mathbf{x}^{\prime}\right)\right\}\right] \\
& =2 \mathbb{E}_{s(\mathbf{w}), b \sim U[0,2 \pi]}\left[\cos \left(\mathbf{w}^{\top} \mathbf{x}+b\right) \cos \left(\mathbf{w}^{\top} \mathbf{x}^{\prime}+b\right)\right] .
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The expectation can be approximated by a Monte Carlo average!
We can reduce the variance of the estimator by generating $M$ samples:

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C\left(\mathbf{x}, \mathbf{x}^{\prime}\right) \approx \frac{2}{M} \sum_{m=1}^{M} \cos \left(\mathbf{w}_{m}^{\top} \mathbf{x}+b_{m}\right) \cos \left(\mathbf{w}_{m}^{\top} \mathbf{x}^{\prime}+b_{m}\right)=\phi(\mathbf{x})^{\top} \phi\left(\mathbf{x}^{\prime}\right)
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For the squared exponential covariance function $s(\mathbf{w})$ is Gaussian!

## Approximate Covariance Function

The covariance matrix can be simply approximated as:

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\boldsymbol{\Sigma} \approx \tilde{\boldsymbol{\Sigma}}=\boldsymbol{\Phi} \Phi^{\top}
$$

and hence $\mathbf{I} \sigma^{2}+\tilde{\boldsymbol{\Sigma}}$ can be inverted with cost $\mathcal{O}\left(N M^{2}\right)$.

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The computational cost is $\mathcal{O}\left(N M^{2}\right)$ !

## Random Features: Illustrative Example

Full GP


## Random Features: Illustrative Example

Random Features GP (M=10)


## Random Features: Illustrative Example

Random Features GP ( $\mathrm{M}=50$ )


## Random Features: Illustrative Example

Random Features GP (M=50)


In regions with no data the approximation may be wiggling a lot!

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- Very simple to implement!
- Equivalent to a neural network with a hidden layer with $M$ units and cosine activations, and a Bayesian linear model in the last layer!


## Full Independent Training Conditional (FITC)

1. Extend model with $M \ll N$ inducing points and outputs at $\overline{\mathbf{X}}$.

$$
p(\mathbf{f}, \mathbf{u})=\mathcal{N}\left(\left.\left[\begin{array}{c}
\mathbf{f} \\
\mathbf{u}
\end{array}\right] \right\rvert\,\left[\begin{array}{l}
\mathbf{0} \\
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2. Introduce conditional independences:

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$$

3. Marginalize $\mathbf{u}$ to obtain an approximate GP prior for $\mathbf{f}$.

$$
p(\mathbf{f})=\int p(\mathbf{f} \mid \mathbf{u}) p(\mathbf{u}) d \mathbf{u}=\prod_{i=1}^{N} p\left(f_{i} \mid \mathbf{u}\right) p(\mathbf{u}) d \mathbf{u}=\mathcal{N}\left(\mathbf{f} \mid 0, \tilde{\Sigma}_{\mathrm{ff}}\right)
$$

where $\tilde{\Sigma}_{\mathrm{ff}}=\mathbf{D}+\mathbf{Q}_{\mathrm{ff}}$ with $\mathbf{D}$ diagonal and $\mathbf{Q}_{\mathrm{ff}}=\boldsymbol{\Sigma}_{\mathrm{fu}} \boldsymbol{\Sigma}_{\mathrm{uu}}^{-1} \boldsymbol{\Sigma}_{\mathrm{uf}}$ of rank $M$.

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5. We make the prediction of $f^{\star}$ at $\mathbf{x}^{\star}$ by considering the approximate GP prior:
$p\left(\mathbf{f}, \mathbf{f}^{\star}\right)=\mathcal{N}\left(\left.\left[\begin{array}{c}\mathbf{f} \\ \mathbf{f}^{\star}\end{array}\right] \right\rvert\,\left[\begin{array}{l}\mathbf{0} \\ \mathbf{0}\end{array}\right],\left[\begin{array}{cc}\tilde{\Sigma}_{\mathrm{ff}} & \mathbf{Q}_{\mathbf{f f}^{\star}} \\ \mathbf{Q}_{\mathbf{f}^{\star f}} & \boldsymbol{\Sigma}_{\mathbf{f}^{\star \mathbf{f}^{\star}}}\end{array}\right]\right)$


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$$
\begin{aligned}
& p\left(\mathbf{y}_{1}, \mathbf{y}_{2}\right)=\mathcal{N}\left(\left[\begin{array}{l}
\mathbf{y}_{1} \\
\mathbf{y}_{2}
\end{array}\right],\left[\begin{array}{l}
\mathbf{a} \\
\mathbf{b}
\end{array}\right],\left[\begin{array}{cc}
\mathbf{A} & \mathbf{C} \\
\mathbf{C}^{\top} & \mathbf{B}
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6. How do we find the location of the inducing points $\overline{\mathbf{X}}$ ?

Simply treat them as prior parameters and maximize the approximate marginal likelihood $p\left(\mathbf{f} \mid \mathbf{0}, \tilde{\Sigma}_{\mathrm{ff}}\right)$ !

## FITC: Illustrative Example

Full GP


## FITC: Illustrative Example

FITC ( $M=10$ )


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The inducing points cover the regions where the function changes!

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- The optimized inducing points spread over the input space where the latent function changes.
- Guaranteed to be exact if $M=N$ and the inducing points are not optimized and located at the training points.
- It can be understood as considering heteroscedastic (input dependent) noise!


## Generalized FITC

Combines FITC with the use of Expectation Propagation to address binary classification problems!

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Approximates with a Gaussian distribution the intractable posterior:

$$
p(\mathbf{f} \mid \mathbf{y})=\frac{\prod_{i=1}^{N} \phi\left(y_{i} f\left(\mathbf{x}_{i}\right)\right) \mathcal{N}(\mathbf{f} \mid \mathbf{0}, \tilde{\boldsymbol{\Sigma}})}{p(\mathbf{y})}
$$

where $\tilde{\boldsymbol{\Sigma}}$ is the approximate FITC covariance matrix.

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Approximates an intractable distribution $p$ by a parametric distribution $q$.

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The exponential family:

$$
q(\mathbf{x})=\exp \left(\boldsymbol{\eta}^{\top} \mathbf{u}(\mathbf{x})-g(\boldsymbol{\eta})\right), \quad g(\boldsymbol{\eta})=\log \int \exp \left(\boldsymbol{\eta}^{\top} \mathbf{u}(\mathbf{x})\right) d \mathbf{x}
$$

where $\boldsymbol{\eta}$ is a vector of natural parameters of $\boldsymbol{q}, \mathbf{u}(\mathbf{x})$ are the sufficient statistics and $g(\boldsymbol{\eta})$ is a $\mathbf{l o g}$ partition function.

## Examples of Distributions in the Exponential Family

Gaussian:

$$
\mathcal{N}\left(x \mid \mu, \sigma^{2}\right)=1 / \sqrt{2 \pi \sigma^{2}} \exp \left\{-\frac{1}{2 \sigma^{2}}(x-\mu)^{2}\right\}
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Exponential form:

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\mathcal{N}\left(x \mid \mu, \sigma^{2}\right)=\exp \left(\boldsymbol{\eta}^{\top} \mathbf{u}(x)-g(\boldsymbol{\eta})\right)
$$

$\boldsymbol{\eta}=\left(\mu / \sigma^{2}, 1.0 / \sigma^{2}\right)^{\top}, \quad \mathbf{u}(x)=\left(x,-0.5 x^{2}\right)^{\top}, \quad g(\boldsymbol{\eta})=-\frac{1}{2} \log \frac{2 \pi}{\eta_{2}}+\frac{\eta_{1}^{2}}{2 \eta_{2}}$.

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Gaussian:

$$
\mathcal{N}\left(x \mid \mu, \sigma^{2}\right)=1 / \sqrt{2 \pi \sigma^{2}} \exp \left\{-\frac{1}{2 \sigma^{2}}(x-\mu)^{2}\right\}
$$

Exponential form:

$$
\mathcal{N}\left(x \mid \mu, \sigma^{2}\right)=\exp \left(\boldsymbol{\eta}^{\top} \mathbf{u}(x)-g(\boldsymbol{\eta})\right)
$$

$\boldsymbol{\eta}=\left(\mu / \sigma^{2}, 1.0 / \sigma^{2}\right)^{\top}, \quad \mathbf{u}(x)=\left(x,-0.5 x^{2}\right)^{\top}, \quad g(\boldsymbol{\eta})=-\frac{1}{2} \log \frac{2 \pi}{\eta_{2}}+\frac{\eta_{1}^{2}}{2 \eta_{2}}$.

Most parametric distributions belong to the exponential family!

## Product and Ratio of Gaussians

Consider these two Gaussian distributions:

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\begin{aligned}
& p_{1}(x)=\frac{1}{\sqrt{2 \pi \sigma_{1}^{2}}} \exp \left\{-\frac{1}{2} \sigma_{1}^{2}\left(x-\mu_{1}\right)^{2}\right\} \\
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## KL-Divergence Minimization

Consider the KL-divergence between $p$ and $q$ ( $q$ in the exponential family):

$$
\mathrm{KL}(p \| q)=-\int p(\mathbf{x}) \log \left\{\frac{q(\mathbf{x})}{p(\mathbf{x})}\right\} d \mathbf{x}=g(\boldsymbol{\eta})-\boldsymbol{\eta}^{\top} \mathbb{E}_{p}[\mathbf{u}(\mathbf{x})]+\text { Const }
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When minimizing with respect to the natural parameters $\boldsymbol{\eta}$ of $q$ :

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$\mathrm{KL}(p \| q)$ is minimized by matching expected sufficient statistics.
If $q$ is Gaussian, then we have to match $\mathbb{E}_{q}[\mathbf{x}]=\mathbb{E}_{p}[\mathbf{x}]$ and

$$
\mathbb{E}_{q}\left[\mathrm{xx}^{\top}\right]=\mathbb{E}_{p}\left[\mathrm{xx}^{\top}\right]
$$

## Joint Approximation

EP approximates this joint distribution by a product of simpler factors:

$$
p(\mathbf{f}, \mathbf{y})=\prod_{i=1}^{N} \phi_{i}\left(y_{i} f\left(\mathbf{x}_{i}\right) \mathcal{N}(\mathbf{f} \mid \mathbf{0}, \tilde{\boldsymbol{\Sigma}})=\prod_{i} t_{i}(\mathbf{f}) \approx \prod_{i} \tilde{t}_{i}(\mathbf{f}),\right.
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where each $\tilde{t}_{i}$ approximates the corresponding $t_{i}$. Each $\tilde{t}_{i}$ must belong to the exponential family but need not be normalized.

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The exponential family is closed under the product and $\prod_{i} \tilde{t}_{i}$ can be easily normalized to compute an approximate distribution:

$$
p(\mathbf{f} \mid \mathbf{y})=\frac{1}{p(\mathbf{y})} \prod_{i} t_{i}(\mathbf{f}) \approx \frac{1}{Z} \prod_{i} \tilde{t}_{i}(\mathbf{f})=q(\mathbf{f})
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Therefore $q$ has the same form as the approximate factors!

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Suppose we wish to refine $\tilde{t}_{j}$. We first remove this factor from the product:

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q^{\bigvee j}(\mathbf{f}) \propto \prod_{i \neq j} \tilde{t}_{i}(\mathbf{f}) \propto q(\mathbf{f}) / \tilde{t}_{j}(\mathbf{f})
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$$

Then, $\tilde{t}_{j}$ is updated to minimize the KL-divergence between:

$$
q_{\text {new }}(\mathbf{f}) \propto \tilde{t}_{j}(\mathbf{f}) q^{\backslash j}(\mathbf{f}), \quad \hat{p}_{j}(\mathbf{f})=\frac{1}{Z_{j}} t_{j}(\mathbf{f}) q^{\backslash j}(\mathbf{f}), \quad Z_{j}=\int t_{j}(\mathbf{f}) q^{\backslash j}(\mathbf{f}) d \mathbf{f}
$$

where $q \backslash j$ is fixed. This ensures that $\tilde{t}_{j}$ is accurate where $q \backslash$ is high.

## Approximate Factors

In practice, $\tilde{t}_{j}$ is found by first minimizing with respect to $q_{\text {new }}$ :

$$
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The refined factor $\tilde{t}_{j}$ is set in practice to be:

$$
\tilde{t}_{j}(\mathbf{f})=Z_{j} \frac{q_{\text {new }}(\mathbf{f})}{q \backslash(\mathbf{f})}, \quad \text { with } \quad \tilde{t}_{j}(\mathbf{f}) q^{\backslash j}(\mathbf{f}) \propto q_{\text {new }}
$$

which ensures that $\tilde{t}_{j}(\mathbf{f}) q{ }^{j}(\mathbf{f})$ and $t_{j}(\mathbf{f}) q{ }^{j}(\mathbf{f})$ integrate the same.

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Several passes are made trough the factors until they converge. The model evidence is approximated by the normalizing constant of $q$.

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The FITC prior results in a total cost of $\mathcal{O}\left(N M^{2}\right)$ !

## Graphical Illustration

Approximates $p(\mathbf{f} \mid \mathbf{y}) \propto t_{0}(\mathbf{f}) \prod_{j=1}^{N} t_{j}(\mathbf{f})$ with $q(\mathbf{f}) \propto t_{0}(\mathbf{f}) \prod_{j=1}^{N} \tilde{t}_{j}(\mathbf{t})$

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$$
p(\mathbf{f} \mid \mathbf{y}) \propto t_{0}(\mathbf{f}) t_{1}(\mathbf{f}) t_{2}(\mathbf{f}) t_{3}(\mathbf{f}), q(\mathbf{f}) \propto t_{0}(\mathbf{f}) \tilde{t}_{1}(\mathbf{f}) \tilde{t}_{2}(\mathbf{f}) \tilde{t}_{3}(\mathbf{f})
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The $\tilde{t}_{j}$ are tuned by minimizing the KL-divergence
$\mathrm{KL}\left[\hat{p}_{j} \| q\right] \quad$ for $j=1, \ldots, N, \quad$ where

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$$

If the exact factor already belongs to the exponential family it needs not be approximated!

## GFITC: Factor Approximation



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The approximate factor is accurate in regions of high posterior probability as indicated by the cavity distribution!

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If EP converges the gradient of $\log Z$ w.r.t. each $\theta_{i}$ is zero, which allows to easily compute the gradients of $\log Z$ !

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\mathbf{f}^{\star} \\
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\end{array}\right] \right\rvert\,\left[\begin{array}{l}
\mathbf{0} \\
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\mathbf{Q}_{\mathbf{f f}} & \Sigma_{\mathrm{ff}}
\end{array}\right]\right)
$$

The conditional $p\left(\mathbf{f}^{\star} \mid \mathbf{f}\right)$ is:

$$
p\left(\mathbf{f}^{\star} \mid \mathbf{f}\right)=\mathcal{N}\left(\mathbf{f}^{\star} \mid \mathbf{Q}_{\mathbf{f}^{\star} \times} \tilde{\Sigma}_{f f}^{-1} \mathbf{f}, \Sigma_{\mathbf{f}^{\star} \neq \star}-\mathbf{Q}_{\mathbf{f}^{\star}} \tilde{\Sigma}_{\mathrm{ff}}^{-1} \mathbf{Q}_{\mathrm{ff}^{\star}}\right)
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## GFITC: Predictions

We want to compute the value of $\mathbf{f}^{\star}$ at a new $\mathbf{x}^{\star}$ :

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p\left(\mathbf{f}^{\star}, \mathbf{f}\right)=\mathcal{N}\left(\left.\left[\begin{array}{c}
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After marginalizing $\mathbf{f}$ w.r.t. $q(\mathbf{f})$, we obtain the predictive distribution:

$$
\begin{aligned}
p\left(\mathbf{f}^{\star} \mid \mathbf{y}\right) & =\int p\left(\mathbf{f}^{\star} \mid \mathbf{f}\right) q(\mathbf{f}) d \mathbf{f} \\
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## GFITC: Illustrative Example

Full GP + EP


## GFITC: Illustrative Example

GFITC (M=10)


## GFITC: Illustrative Example

GFITC (M=10)


The inducing points spread across the input space!

## Variational Free Energy

Previous methods approximate the GP prior using a low rank approximation of $\Sigma$, resulting in a $\operatorname{cost} \mathcal{O}\left(N M^{2}\right)$.

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- Variational inference is used to tune $q$.

Since the GP prior is not changed it tends to perform better than the previous methods!

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Adjust the parameters of $q$ to match $p$ by minimizing $\mathrm{KL}(q \mid p) \geq 0$.

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$\mathrm{KL}(q \mid p)$ depends on $p$, which is assumed to be intractable!
Let the target be $p(\mathbf{f} \mid \mathbf{y})$. Consider the decomposition of $p(\mathbf{y})$ :

$$
\log p(\mathbf{y})=\mathcal{L}(q)+\mathrm{KL}(q \mid p)
$$

where

$$
\mathcal{L}(q)=\int q(\mathbf{f}) \log \frac{p(\mathbf{f}, \mathbf{y})}{q(\mathbf{f})} d \mathbf{f}, \quad \mathrm{KL}(q \mid p)=\int q(\mathbf{f}) \log \frac{q(\mathbf{f})}{p(\mathbf{f} \mid \mathbf{y})} d \mathbf{f} .
$$

## Decomposition of the Marginal Likelihood



## Decomposition of the Marginal Likelihood


$\mathcal{L}(q)$ can be used to approximate $\log p(\mathbf{y})$ if $\mathrm{KL}(q \mid p)$ is small!

## Variational Free Energy (VFE)

Lower bound the log-likelihood:
$\log p(\mathbf{y} \mid \theta)=\log \int p(\mathbf{y}, \mathbf{f}, \mathbf{u} \mid \theta) d \mathbf{f} d \mathbf{u}$


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& \left.\left.\mathcal{L}(q, \theta)=\int q(\mathbf{f}, \mathbf{u}) \log \frac{p(\mathbf{y}, \mathbf{f}, \mathbf{u} \mid \theta)}{q(\mathbf{f}, \mathbf{u})} d \mathbf{f} d \mathbf{u} \right\rvert\, \theta\right)=\log p(\mathbf{y} \mid \theta)-\operatorname{KL}[q(\mathbf{f}, \mathbf{u}) \mid p(\mathbf{f}, \mathbf{u} \mid \mathbf{y})]
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KL $\equiv$ Kullback-Leibler divergence

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\end{aligned}
$$


$K L \equiv$ Kullback-Leibler divergence
By maximizing $\mathcal{L}(q, \theta)$ w.r.t $q$ we are enforcing that $q(\mathbf{f}, \mathbf{u})$ looks similar to $p(\mathbf{f}, \mathbf{u} \mid \mathbf{y})$ in terms of the KL!

## Variational Free Energy (VFE)

Consider the following approximate distribution:

$$
q(\mathbf{f}, \mathbf{u})=p(\mathbf{f} \mid \mathbf{u}) \quad q(\mathbf{u})=p(\mathbf{f} \mid \mathbf{u}) \mathcal{N}(\mathbf{u} \mid \mathbf{m}, \mathbf{S})
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- Fixed
- Tunable-


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- Fixed
- Tunable


The inducing points are now parameters of the approx. dist. $q$ !

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Plugging $q(\mathbf{f}, \mathbf{u})$ into the lower bound we have:

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& =\int p(\mathbf{f} \mid \mathbf{u}) q(\mathbf{u}) \log \frac{p(\mathbf{y} \mid \mathbf{f}, \theta) p(\mathbf{f} \mid \mathbf{u}) p(\mathbf{u})}{p(\mathbf{f} \mid \mathbf{u}) q(\mathbf{u})} d \mathbf{f} d \mathbf{u}
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\end{array} \\
\mathcal{L}(q, \theta)=\mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y} \mid \mathbf{f}, \theta)]-\operatorname{KL}[q(\mathbf{u}) \mid p(\mathbf{u})]
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& \mathcal{L}(q, \theta)=\mathbb{E}_{q(\mathbf{f})}[\log p(\mathbf{y} \mid \mathbf{f}, \theta)]-\mathbf{K L}[q(\mathbf{u}) \mid p(\mathbf{u})] \\
& \text { - Mean squared prediction error } \\
& \text { - KL between Gaussians }
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- No change in the model is made and the cost is in $\mathcal{O}\left(M^{2} N\right)$ !


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- No change in the model is made and the cost is in $\mathcal{O}\left(M^{2} N\right)$ !
- Predictions are made using $p\left(\mathbf{f}^{\star} \mid \mathbf{u}\right) q(\mathbf{u})$ marginalizing out $\mathbf{u}$.


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We want to compute the value of $\mathbf{f}^{\star}$ at a new $\mathbf{x}^{\star}$ :

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\end{aligned}
$$

## VFE: Illustrative Example

Full GP


## VFE: Illustrative Example

VFE ( $M=10$ )


## VFE: Illustrative Example

$\operatorname{VFE}(\mathrm{M}=10)$


The inducing points cover the regions where the function changes!

## VFE: Illustrative Classification Example

Full GP + EP


## VFE: Illustrative Classification Example

VFE (M=10)


## VFE: Illustrative Classification Example

VFE (M=10)


The inducing points spread across the input space!

## FITC vs. VFE

Two approaches:

## FITC vs. VFE

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- FITC: optimize the marginal likelihood of an approximate GP model.
- VFE: maximize fidelity to the original exact GP.


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FITC


VFE


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Two approaches:

- FITC: optimize the marginal likelihood of an approximate GP model.
- VFE: maximize fidelity to the original exact GP.

FITC


VFE


- FITC: less local optima and easier to optimize, also less accurate.
- VFE: more accurate, more local optima, more difficult to optimize.
(Bui et al., 2017) (Bauer et al., 2016)


## Whitened Parameterization for VFE

Alternative VFE objective expected to be easier to optimize!

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Instead of making inference about $\mathbf{u}$, the whitened VFE objective makes inference about:

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\mathbf{e} \text { such that } \mathbf{u}=\mathbf{L e}, \quad \mathbf{e} \sim \mathcal{N}(\mathbf{0}, \mathbf{I})
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with $\mathbf{u}$ the latent process values at the inducing points and $\mathbf{L}^{\top} \mathbf{L}=\boldsymbol{\Sigma}_{\mathbf{u u}}$.

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with $\mathbf{u}$ the latent process values at the inducing points and $\mathbf{L}^{\top} \mathbf{L}=\boldsymbol{\Sigma}_{\mathbf{u u}}$.
The VFE objective becomes:

$$
\sum_{i=1}^{N} \mathbb{E}_{q(\mathbf{e}) p\left(f\left(\mathbf{x}_{i}\right) \mid \mathbf{e}\right)}\left[\log p\left(y_{i} \mid f\left(\mathbf{x}_{i}\right)\right)\right]-\mathrm{KL}(q(\mathbf{e}) \mid \mathcal{N}(\mathbf{0}, \mathbf{I}))
$$

with $p\left(f\left(\mathbf{x}_{i}\right) \mid \mathbf{e}\right)$ using the covariances between $f\left(\mathbf{x}_{i}\right)$ and $\mathbf{e}$.

## Whitened Parameterization: Illustrative Example



## Whitened Parameterization: Illustrative Example



Whitening significantly improves convergence!

## Natural Gradient Ascent

Gradient ascent moves in the direction of the gradient $\nabla_{\xi} \mathcal{L}(\xi)$.

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Gradient ascent moves in the direction of the gradient $\nabla_{\xi} \mathcal{L}(\xi)$.

Formally:

$$
\nabla_{\boldsymbol{\xi}} \mathcal{L}(\boldsymbol{\xi}) \propto \lim _{\epsilon \rightarrow 0} \frac{1}{\epsilon} \arg \max \operatorname{ds.t.|\mathbf {d}\| \leq \epsilon } \mathcal{L}(\boldsymbol{\xi}+\epsilon \mathbf{d})
$$

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Formally:

$$
\nabla_{\boldsymbol{\xi}} \mathcal{L}(\boldsymbol{\xi}) \propto \lim _{\epsilon \rightarrow 0} \frac{1}{\epsilon} \arg \max \operatorname{ds.t.|\mathbf {d}\| \leq \epsilon } \mathcal{L}(\boldsymbol{\xi}+\epsilon \mathbf{d})
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The steepest ascent direction picks $\mathbf{d}$ in the $\epsilon$-vicinity of $\boldsymbol{\xi}$ (measured by the Euclidean norm) that maximizes $\mathcal{L}(\cdot)$.

## Natural Gradient Ascent

Gradient ascent moves in the direction of the gradient $\nabla_{\xi} \mathcal{L}(\xi)$.

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If $\boldsymbol{\xi}$ represents the parameters of probability distributions, the Euclidean norm may be problematic!

## Illustration with Two Gaussians



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The Euclidean distance between parameters is 4 in both cases!

## Illustration with Two Gaussians



A better alternative is the KL-divergence between distributions!

## Natural Gradient Ascent

Considers the KL-divergence as a norm:
with $\mathbf{F}_{\boldsymbol{\xi}}$ the Fisher information of $q$ :

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\mathbf{F}_{\xi}=-\mathbb{E}_{q(\mathbf{u} \mid \boldsymbol{\xi})}\left[\nabla_{\xi}^{2} \log q(\mathbf{u} \mid \boldsymbol{\xi})\right]
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\mathbf{F}_{\boldsymbol{\eta}}=\frac{\partial \boldsymbol{\theta}}{\partial \boldsymbol{\eta}}, \quad \quad \mathbf{F}_{\boldsymbol{\xi}}=\left(\frac{\partial \boldsymbol{\eta}}{\partial \boldsymbol{\xi}}\right)^{\top} \frac{\partial \boldsymbol{\theta}}{\partial \boldsymbol{\eta}} \frac{\partial \boldsymbol{\eta}}{\partial \boldsymbol{\xi}}
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Thus,

$$
\nabla_{\boldsymbol{\xi}} \mathcal{L}(\boldsymbol{\xi}) \mathbf{F}_{\xi}^{-1}=\frac{\partial \mathcal{L}}{\partial \boldsymbol{\theta}}\left(\frac{\partial \boldsymbol{\xi}}{\partial \boldsymbol{\eta}}\right)^{\top}
$$

## Natural Gradient Ascent


(Salimbeni et al., 2018)

## Natural Gradient: Illustrative Example



## Natural Gradient: Illustrative Example



The natural gradient achieves a faster convergence!

## GPs for Big Data

Can we further improve the computational cost in $\mathcal{O}\left(N M^{2}\right)$ ?

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The training cost goes down to $\mathcal{O}\left(M^{3}\right)$ which allows to address datasets with millions of instances!
(Hensman et al., 2013)

## GPs for Big Data



## GPs for Big Data



To converge to a local neighborhood of the optimum stochastic methods require an estimate of the gradient which can be very cheap!

## GPs for Big Data



Training Time in Seconds in a log10 Scale
(Hernández-Lobato, 2015)

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- It allows for minibatch training which reduces the cost to $\mathcal{O}\left(M^{3}\right)$.
- The objective is prone to local optima and difficult to optimize.


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- The best performing method seems to be the VFE method since it does not modify the prior.
- Some methods allow for stochastic optimization and mini-batch training that further reduce the cost to $\mathcal{O}\left(M^{3}\right)$.


## References

- Williams, C., \& Seeger, M. (2000). Using the Nyström method to speed up kernel machines. Advances in neural information processing systems, 13.
- Snelson, E., \& Ghahramani, Z. (2005). Sparse Gaussian processes using pseudo-inputs. Advances in neural information processing systems, 18.
- Rahimi, A., \& Recht, B. (2007). Random features for large-scale kernel machines. Advances in neural information processing systems, 20.
- Naish-Guzman, A., \& Holden, S. (2007). The generalized FITC approximation. Advances in neural information processing systems, 20.
- Hernández-Lobato, D., \& Hernández-Lobato, J. M. (2016, May). Scalable Gaussian process classification via expectation propagation. In Artificial Intelligence and Statistics (pp. 168-176).
- Hensman, J., Fusi, N., \& Lawrence, N. D. (2013). Gaussian processes for big data. Proceedings of the Twenty-Ninth Conference on Uncertainty in Artificial Intelligence.
- Salimbeni, H., Eleftheriadis, S., \& Hensman, J. (2018, March). Natural gradients in practice: Non-conjugate variational inference in Gaussian process models. In International Conference on Artificial Intelligence and Statistics (pp. 689-697).
- Wu, L., Miller, A., Anderson, L., Pleiss, G., Blei, D., \& Cunningham, J. (2021). Hierarchical inducing point gaussian process for inter-domian observations. International Conference on Artificial Intelligence and Statistics (pp. 2926-2934).

